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Conductance-based dynamic causal modeling: A mathematical review of its application to cross-power spectral densities



(Marreiros et al., 2010b). First proposed in Friston et al. (2003), DCM was initially developed for functional magnetic resonance imaging

(fMRI) data. In DCM for fMRI, the neuronal model is simple, assum-

ing a single state variable per region and modeling neuronal popula-

tion dynamics by using a bilinear or 2nd-order Taylor approximation

(Stephan et al., 2008). This relatively simple model is used to estimate

how specific brain regions interact with each other through directed

synaptic interactions (effective connectivity) and to evaluate whether

and how this connectivity is influenced by experimentally controlled

factors and/or disease processes. David et al. (2006a) extended this mod-

eling technique to event-related responses (ERPs) measured with elec-

troencephalography (EEG) or magnetoencephalography (MEG). In this

seminal work, DCM of ERPs was cast in terms of a neural mass model

based on the Jansen-Rit model (Jansen and Rit, 1995) which represents

a more sophisticated model of neuronal dynamics as compared to the

relatively abstract descriptions in DCM for fMRI. This biological real-

ism is afforded by the richer temporal information contained in electro-

physiological measurements (on the order of milliseconds), as compared

to the coarse nature of the BOLD response (on the order of seconds)

(Huettel et al., 2014).

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ABSTRACT

Dynamic Causal Modeling (DCM) is a Bayesian framework for inferring on hidden (latent) neuronal states, based on measurements of brain activity. Since its introduction in 2003 for functional magnetic resonance imaging data, DCM has been extended to electrophysiological data, and several variants have been developed. Their biophysically motivated formulations make these models promising candidates for providing a mechanistic understanding of human brain dynamics, both in health and disease. However, due to their complexity and reliance on concepts from several fields, fully understanding the mathematical and conceptual basis behind certain variants of DCM can be challenging. At the same time, a solid theoretical knowledge of the models is crucial to avoid pitfalls in the application of these models and interpretation of their results. In this paper, we focus on one of the most advanced formulations of DCM, *i.e.* conductance-based DCM for cross-spectral densities, whose components are described across multiple technical papers. The aim of the present article is to provide an accessible exposition of the mathematical background, together with an illustration of the model's behavior. To this end, we include step-by-step derivations of the model equations, point to important aspects in the software implementation of those models, and use simulations to provide an intuitive understanding of the type of responses that can be generated and the role that specific parameters play in the model. Furthermore, all code utilized for our simulations is made publicly available alongside the manuscript to allow readers an easy hands-on experience with conductance-based DCM.

1. Introduction

Dynamic Causal Modeling (DCM) is a framework to construct generative models describing how putative neural mechanisms give rise to neurophysiological data. By inverting these models, using generic Bayesian techniques, it is possible to infer upon hidden (unobserved) neuronal states from measured data (for reviews, see Daunizeau et al. 2011, Friston et al. 2013, Kiebel et al. 2008b, Stephan et al. 2009). As a modeling technique, DCM has been used to study physiological processes in the healthy human brain (for examples, see Garrido et al. 2008, Grefkes et al. 2008, Summerfield et al. 2006). In addition, it has also been increasingly employed in the fields of Computational Psychiatry (Adams et al., 2020; Breakspear et al., 2015; Brodersen et al., 2014; Frässle et al., 2020, 2018; Schmidt et al., 2013) and Computational Neurology (Brodersen et al., 2011; Cooray et al., 2015; Gilbert et al., 2016; Rosch et al., 2018; Shaw et al., 2019; Symmonds et al., 2018), with the hope of not only providing mechanistic insights into pathophysiology, but also of developing the model into a useful clinical tool (i.e., a "computational assay", Stephan et al., 2006; Stephan and Mathys, 2014).

 $DCMs^1$ are formulated using ordinary or stochastic differential equations which describe the dynamics of neural (hidden) states over time

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¹ In this paper, we used the acronym "DCM" both to refer to the modeling approach (dynamic causal modeling) and to its instantiation (dynamic causal model).



Fig. 1. General structure of DCMs. A DCM for electrophysiological data comprises two parts: (1) the **neuronal model**, which delineates the dynamics $f(x(t), u(t), \theta)$ of the hidden states x(t), as well as ensuing neuronal population activity, and (2) the **observation model**, which describes how source activity h(x(t)) propagates through surrounding tissues (brain, skull, scalp) in order to generate the data y(t) registered at the level of the sensors. u(t) are the inputs to a neuronal population, θ represents the model parameters and Γ_x and ϵ are stochastic noise terms. Figure created with Biorender.com.

Marreiros et al. (2009) introduced an important variant of DCM for electrophysiological data, describing a conductance-based model based on the Morris–Lecar model (Morris and Lecar, 1981). This conductancebased DCM (cbDCM) contains explicit representations of ionotropic receptors with distinct time constants, namely the AMPA, NMDA and GABA_A receptors. Given the central role of these ionotropic receptors in many psychiatric and neurological disease processes and given that drugs targeting NMDA and GABA_A receptors exist, this mechanistically fine-grained formulation is of considerable interest for studying neural circuits and their alterations in disease and under pharmacological interventions. For example, cbDCM has been applied to data from patients with monogenic channelopathies (Gilbert et al., 2016) and NMDA receptor antibody encephalitis (Symmonds et al., 2018) as well as to data from pharmacological studies in healthy volunteers and animals (Adams et al., 2020a; Moran et al., 2015, 2011b).

Unfortunately, the literature introducing cbDCM and its mathematical foundations is distributed over several papers and is not easily accessible for the average neuroscientist or clinician. This tutorial-style paper offers a didactic treatment of the model, focusing on the underlying theory and mathematical derivations. These derivations are here expanded beyond what is presented in the original papers, in order to provide a more detailed, step-by-step description of the model. In addition, several footnotes will be added along the way, so as to offer extra hints without breaking the flow of the manuscript. Furthermore, we will discuss practical aspects related to the inversion of DCMs using Statistical Parametric Mapping (SPM), a freely available open-source and widely used software package written in MATLAB. Notably, this paper is not meant to provide an exhaustive review of the existing literature on variants of DCM for EEG/MEG. Comprehensive reviews on this topic as well as general practice recommendations can be found in other publications (Kiebel et al., 2008b; Moran et al., 2013; Stephan et al., 2009). Finally, this paper will not cover the topics of Bayesian inference or model inversion in detail. More in-depth accounts of these concepts can be found elsewhere (Bishop, 2006; Friston et al., 2006).

Generally, a DCM for electrophysiological data comprises two parts (Fig. 1): the **neuronal model**, which delineates the intra- and interneuronal source² dynamics, and the **observation model**, which de-

scribes how source activity propagates through surrounding tissues (brain, skull, scalp) in order to generate the data registered at the level of the sensors (Kiebel et al., 2008c). In this paper, for conceptual clarity, we will first review how to characterize and distinguish the existing variants of DCM for EEG/MEG and local field potentials (LFP). In doing so, we will provide insights as to which model could be used, depending on the research question considered. Secondly, we offer an introduction to conductance-based models. We outline the neuronal model of cbDCM, moving from the single-neuron level to the population level, to the level of the cortical column, before finally describing connectivity amongst sources. In addition, we give an account of the observation model for cross-spectral densities. In parallel, we review the literature where this model was introduced and described. However, since these seminal papers were published, the implementation of these models in SPM has undergone several refinements and modifications. Therefore, we also review some of the newer aspects of SPM that are relevant for the user.

The paper assumes that the reader has basic knowledge of neuroanatomy, neurophysiology, Bayesian statistics, and signal processing techniques such as the Fourier transform and convolutions. A list of important concepts, along with their definitions, is provided in Table 1.

2. DCM variants

DCMs for electrophysiological data come in two different flavors: neurophysiological and phenomenological. Phenomenological models include DCM for induced responses and DCM for phase coupling (Penny et al., 2009; Yeldesbay et al., 2019), and are characterized by an evolution function of the neuronal states that is not closely related to the underlying neurophysiology. These models will not be addressed in this paper; however, for details, please consult the relevant literature (Chen et al., 2008; Penny et al., 2009). The focus of the present paper will therefore be on neurophysiologically-informed DCM variants, which can again be divided into several categories depending on several dimensions (compare Fig. 2):

- How the cortical column is described We distinguish "convolution-based" models, which primarily consider the cortical column, from "conductance-based" variants, which start by modeling a single cell's electrophysiological properties (Moran et al., 2013). In addition, there are different ways of modeling a cortical column. The first DCMs for electrophysiological data were based on the Jansen-Rit model (Jansen and Rit, 1995) and included 3 neuronal populations (excitatory pyramidal cells, excitatory spiny stellate cells, and inhibitory interneurons), whilst the more recent "canonicalmicrocircuit" variant takes 4 populations into account (superficial and deep pyramidal cells, spiny stellate cells and inhibitory interneurons) (Moran et al., 2013).
- How the hidden states of the neuronal populations are modeled – *i.e.*, whether a population's density is summarized by a single number (first-order statistic), as is the case for neural-mass models (NMM); or whether higher-order statistics are also taken into account, as is the case for mean-field models (MFM). If one makes the neuronal states not only a function of time (as in the NMM and MFM), but also a function of space, one obtains a neural-field model (NFM).
- Whether or not there is an exogenous/experimental input Three types of neuronal activity can be captured: (1) Event-related potentials (ERP), which correspond to the phase-locked response of a neural system to a stimulus and can be characterised by averaging in the time domain (David et al., 2006b). (2) Induced responses, which constitute changes in neuronal oscillations that appear after a stimulus, but are not phase-locked to this stimulus; they can be obtained by averaging in the frequency domain (David et al., 2006b; Tallon-Baudry and Bertrand, 1999). As explained before, DCM for induced responses (Chen et al., 2008) will not be discussed in detail in this paper. (3) "Resting-state" or "spontaneous" activity, where no exoge-

 $^{^2}$ In this context, a "source" refers to a neuronal population (or region) of the brain. This terminology is common because activity in a neuronal population constitutes the *source* of signal.

Table 1

Important concepts and definitions.

| Concept | Definition | | | |
|--------------------------|---|--|--|--|
| Statistic | Index of some attribute of given data (e.g., sample mean). "Moment" is sometimes used interchangeably with "statistic", where the first moment corresponds to the mean and the second moment to the covariance. | | | |
| Time series | "A series of values of a quantity obtained at successive times, often with equal intervals between them" (Oxford Dictionary, 2020a). | | | |
| Resting-state | A condition of unconstrained cognition in which there is no exogenous, experimental input and brain activity is therefore "spontaneous". | | | |
| Sensor | "Device which detects or measures a physical property and records, indicates, or otherwise responds to it" (Oxford Dictionary, 2020b). In the context of this paper, it is used interchangeably with "channel" or "electrode". | | | |
| Frequency band | "Range of frequencies [] between two limits" (Oxford Dictionary, 2020c). The main frequency bands used in EEG analysis include the delta (<4 Hz), theta (4–7 Hz), alpha (8–12 Hz), beta (13–29 Hz) and gamma (30–79) bands (Mecarelli, 2019). | | | |
| Hidden state | Unobserved quantity. "Hidden state" is used interchangeably with "latent state". In deterministic forms of DCMs (where trajectories of states are fully determined by the values of the model's parameters and [known] inputs), we wish to model the neuronal hidden states by inferring upon the model parameters using observable data. | | | |
| Density | In the context of this publication, the term "density" (e.g., as in "ensemble density") refers to the probability density function of a continuous random variable. | | | |
| Generative model | A probabilistic model that describes the putative process by which data were generated. Mathematically speaking, generative models specify the joint probability density (product of likelihood and prior) over model parameters and measured data. By sampling from the prior, it is possible to generate synthetic data points (Bishop, 2006). | | | |
| Forward model | Mapping from hidden states (e.g., neuronal activity) to observed data (e.g., EEG measurements). "Forward model" will be used interchangeably with "observation model" (Moran et al., 2013). A neuronal model and an observation model together make up a DCM. | | | |
| Model inversion | In a Bayesian setting, the process by which a generative model is used to compute the posterior distribution of the model parameters. | | | |
| Mean field model (MFM) | A population (ensemble) of spiking neurons can be modeled using a population density function which describes the probabilistic evolution of the population response over time. Such a mean field model (MFM) is described by the dynamics of the moments of the population density function (Moran et al., 2013). MFMs have been used for more than half a century and are defined using concepts from statistical physics (refer to the Fokker-Planck equation below) (Beurle, 1956; Deco et al., 2008). Since EEG/MEG data reflect the activity of populations of neurons, MFMs are well suited for these data modalities. | | | |
| Neural mass model (NMM) | If one only considers the first moment (<i>i.e.</i> , the mean) of a MFM, one obtains a neural mass model (NMM) (Deco et al., 2008). NMMs are therefore a special case of MFMs (Moran et al., 2013). | | | |
| Neural field model (NFM) | NMMs and MFMs consider the evolution of the neuronal states only over time. However, this evolution can also be modeled over space (<i>e.g.</i> , across the cortical sheet). Models involving differential operators with both temporal and spatial terms are called neural field models (Deco et al., 2008). | | | |
| Fokker-Planck equation | The Fokker-Planck equation is a partial differential equation from statistical physics that allows one to describe, using a flow-diffusion process, the evolution of a probability density function of an ensemble of individual components (<i>e.g.</i> , neurons) over time. The advantage offered by this approach is that it permits modeling of ensemble density dynamics in a deterministic manner, even if the dynamics of individual components are stochastic. | | | |

nous (experimental) input is present. In the latter case, the output is summarized in the frequency domain, *i.e.*, we obtain resting-state spectral oscillations.

How source data are estimated from the observed sensor data

 There are two main approaches to this (Kiebel et al., 2008a): (1) point-source models explain the data using a small number of so-called equivalent current dipoles (ECDs) to represent each source, while (2) imaging models use a dense set of dipoles distributed over the cortical sheet (Daunizeau et al., 2006). In addition, the type of data modeled, *i.e.*, EEG data, MEG data, or local field potentials (LFP), also influence the specific form of the forward model (Table 1), in particular, the form of the so-called "lead field matrix" (please refer to Section 8 for a more detailed discussion on these topics).

Importantly, the defining aspects considered above can be combined. For example, it is possible to apply neural-mass or mean-field formulations to convolution or conductance-based models (Moran et al., 2013). Choosing an adequate DCM variant involves careful consideration of the data and problem at hand. This is because different (clinical) questions might be more naturally addressed by some variants than others. For instance, in an autoimmune disorder called anti-NMDA receptor encephalitis, auto-immune antibodies selectively target the NMDA receptor, leading to receptor hypofunction and a variety of severe psychiatric and neurological conditions, including psychosis and epilepsy. Hence, in order to test disease-related hypotheses about NMDA receptor function directly, it may be advisable to use a DCM that can represent this receptor explicitly, such as cbDCM. This was precisely the modeling approach followed by Symmonds et al. (2018) who used cbDCM to model EEG data from anti-NMDA receptor encephalitis patients and controls. Hence, the choice of the optimal DCM variant has to be tailored to the specific hypothesis about (disease-relevant) processes of interest.

In the following, we define what constitutes a conductance-based model.

3. Conductance-based models

The electrical properties of neurons can be described by the membrane conductance associated with different ions. Hodgkin and Huxley (1952) famously used this formalism to describe how sodium and potassium currents can generate action potentials in the giant axon of the squid. It has been extended to include other ions (see Morris and Lecar, 1981) and can accommodate active, neurotransmitter-mediated ion flow, as well as leaky ion channels and externally applied current (Koch, 1999). Conductance-based models are hence models of excitable cells (*e.g.*, neurons) that represent ion channels through their conductance (Skinner, 2006). According to Dayan and Abbott, these models have been shown to "*reproduce the rich and complex dynamics of real neurons quite accurately*" (Dayan and Abbott, 2001).

In the following sections, we will describe the conductance-based neuronal model implemented in DCM for electrophysiological data. In our description, we will move from the single-neuron level, to the level of the cortical column, and finally describe macroscopic source connectivity.



Fig. 2. DCM variants, along with examples of relevant publications (Brown and Friston, 2012; David et al., 2006; Friston et al., 2012; Garrido et al., 2007; Gilbert et al., 2016; Kiebel et al., 2006; Marreiros et al., 2009, 2010a; Moran et al., 2007a, 2011a; Pinotsis et al., 2012, 2013). Models of induced responses are not considered. EEG: electroencephalography; LFP: local field potential; MEG: magnetoencephalography; MFM: mean field model; NFM: neural field model; NMM: neural mass model. Figure created with Biorender.com.

4. Modeling a single neuron

Conductance-based models describe how ions flow into and out of excitable cells via a parallel resistor-capacitor (RC) circuit. Using these so-called equivalent circuit models, it is possible to model the dynamics associated with specific ions or specific ligand-gated ion channels (also referred to as ionotropic receptors). Let us consider a neuron with two glutamate receptors, AMPA and NMDA, and the GABA_A receptor³ (see Fig. 3). Note that both the AMPA and NMDA receptors are permeable to multiple cations, not just sodium. A leak current *L* is also included, to model the effect of other passive ion channels on the cell's resting membrane potential.

The equation of motion for the membrane potential *V* can be derived by making use of two fundamental laws of physics. Specifically, we make use of Kirchhoff's current law (Koch, 1999):

$$u = I_c + \sum_k I_k \tag{1}$$

where *u* is the injected current, I_c the capacitive current, and I_k is the resistive current associated with the *k*-th channel. Rearranging this equation yields:

$$I_c = -\sum_k I_k + u \tag{2}$$

We further use the current-voltage relation of a capacitor to reexpress the capacitive current I_c :

$$I_c = C \cdot \frac{dV}{dt} \tag{3}$$

 $^{^3\,}$ AMPA, NMDA and GABA_A receptors are associated with distinct time constants, a condition necessary for identifying their relative contribution to the measured potentials.



Fig. 4. (Left) Voltage-dependent block of the NMDA receptor pore by Mg^{2+} . Note how channel opening necessitates neurotransmitter binding as well as depolarization of the cell (figure adapted from Purves et al., 2018). (**Right**) $m(V)(V - V_{NMDA})$ represented as a function of membrane voltage. The black line depicts the behavior of $m(V)(V - V_{NMDA})$ for the parameter value currently used in SPM12. Figure created with Biorender.com.

In addition, we apply Ohm's law to express the resistive current associated with each channel *k*:

$$I_k = \frac{V - V_k}{R_k} = g_k \left(V - V_k \right) \tag{4}$$

Here, V_k represents the reversal potential for channel k and the conductance $g_k = 1/R_k$ is given by the inverse resistance R_k . By combining Eqs. (2)–(4), we can relate the membrane capacitive current to the cellular ionic currents.

$$C\dot{V} = \left(\sum_{k} g_{k}(V_{k} - V)\right) + u$$

where : $k \in L$, AMPA, GABA (5)

Note, however, that Eq. (5) does not yet contain a term that explicitly represents the NMDA receptor. This requires a slight extension because of the so-called voltage-dependent magnesium (Mg^{2+}) block of the NMDA receptor. Indeed, at hyperpolarized potentials, magnesium is present within the channel pore, effectively blocking it (Fig. 4, left). Depolarization of the membrane potential pushes Mg^{2+} out of the pore, allowing current to flow through the channel (Purves et al., 2018).

This Mg²⁺ nonlinearity is described as follows (Jahr and Stevens, 1990; Koch, 1999; Moran et al., 2011a):

$$m(V) = \frac{1}{1 + 0.2 \cdot exp(-\alpha_{NMDA} \cdot V)}$$
(6)

where α_{NMDA} represents the magnesium block parameter⁴ (Fig. 4, right).

By adding a term that represents the NMDA receptor, as well as a noise term Γ_V , the final expression for the equation of motion for the membrane potential *V* is obtained (Moran et al., 2011a):

$$C\dot{V} = \left(\sum_{k} g_{k} (V_{k} - V)\right) + g_{NMDA} m(V) (V_{NMDA} - V) + u + \Gamma_{V}$$

$$k \in L, \text{ AMPA, GABA}$$
(7)

Where:

- *C* is the membrane capacitance;
- *V* the membrane potential;
- \dot{V} the time derivative of *V*;
- *g_k* represents the conductance for channel *k*;
- *V_k* is the reversal potential for channel *k*;
- *u* is the applied input current. For cells which receive no external input: *u* = 0;
- Γ_V is a stochastic term which models Gaussian noise;
- L denotes the passive leak current channels.

We hence have our first set of differential equations, which model the change in membrane **voltage** of a single neuron over time. Now, we turn to the channel conductances. The leak channel conductance g_L is assumed to be fixed. Thus, we define the subsequent equation only for

⁴ In the current implementation of SPM12: α_{NMDA} is fixed and set to 0.06 (see mg switch.m).

$k \in \{AMPA, GABA_A, NMDA\}$ (Moran et al., 2011a):

$$\dot{g}_k = \kappa_k (\gamma_{aff} \cdot \sigma_{aff} - g_k) + \Gamma_g \tag{8}$$

where κ_k is the inverse time constant for the *k*th receptor, σ_{aff} models the firing from afferent neurons weighted by γ_{aff} , which represents a coupling parameter (*i.e.*, represents connection strength). Γ_g represents the stochastic component, again modeling Gaussian noise.

In short, the model entails **four hidden states**: the membrane voltage *V* and the conductance g_k for all three receptors: AMPA, GABA_A and NMDA.

5. Modeling a population of neurons

Let us start by examining the equations of motion for the hidden states *x* from the previous section Eqs. (7) and ((8)). Note that they are stochastic differential equations of the form (Fig. 1):

$$\dot{x} = f(x) + \Gamma_{x} \tag{9}$$

This form corresponds to a Langevin equation (Haken, 1983), where *x* denotes a vector of neuronal state variables of interest, f(x) represents the deterministic part of the equation, and Γ_x is the stochastic component. Neurobiologically, this stochastic component represents different sources of noise. An example at the single-neuron level is *thermal noise, i.e.,* thermal fluctuations which trigger spontaneous conformational changes in proteins that are part of the ion channels, some of which lead to opening or closing of the channel (*i.e., channel noise*).⁵ For an extensive review on the sources of noise in the brain, please see other previous publications (Destexhe and Rudolph-Lilith, 2012; Faisal et al., 2008).

At this point, we are still at the level of single neurons. However, techniques for measuring human brain activity noninvasively, such as EEG or MEG, cannot resolve the behavior of single neurons. Instead, they provide measurements of the activity of large populations (or ensembles) of neurons. When moving to models of neuronal population dynamics, capturing the stochastic dynamics of each neuron by Langevin equations would be too complex and computationally expensive. Instead, one can formulate a partial differential equation that describes the temporal evolution of the probability density function of *x*, now seen as a random variable. This is called the Fokker-Planck equation (Haken, 1983; Kadanoff, 2000) (Table 1): it is a deterministic equation that captures the dynamics of a probability density of neuronal population states (ensemble density). In other words, each neuron is treated as a point in state space, the ensemble density describes their probability density over state space, and the Fokker-Planck equation describes how this ensemble density changes over time (Deco et al., 2008).

If we assume that the ensemble density can be represented by a Gaussian distribution $q(x) = \mathcal{N}(\mu, \Sigma)$ (the so-called Laplace approximation), these equations can be reformulated as ordinary differential equations of the sufficient statistics of the population's density. Of course, this assumption might entail some loss of information, if the ensemble density is not well represented by a Gaussian distribution. In any case, under the Laplace approximation, for each hidden state *l* in the *j*-th neuronal population, the equations that describe the dynamics of the sufficient statistics are (Marreiros et al., 2009; Moran et al., 2011a):

$$q(x) = \mathcal{N}(\mu, \Sigma) \tag{10}$$

$$\dot{\mu}_{l}^{(j)} = f_{l}^{(j)}(\mu) + \frac{1}{2}Tr\left(\Sigma^{(j)}\frac{\partial^{2}f_{l}^{(j)}}{\partial x^{2}}\right)$$
(11)

$$\dot{\Sigma}^{(j)} = \frac{\partial f^{(j)}}{\partial x} \Sigma^{(j)} + \Sigma^{(j)} \frac{\partial f^{(j)T}}{\partial x} + D^{(j)} + D^{(j)T}$$
(12)

With Eqs. (10)–(12), we have thus defined a mean-field model (MFM), where, $f_l^{(j)}$ is the deterministic part of the *j*-th neuronal population's *l*-th hidden state value, and *D* represents the so-called diffusion coefficient. From Eq. (11), we can see that in this mean-field formulation the covariance matrix $\Sigma^{(j)}$ affects the dynamics of the mean of the ensemble density. By contrast, in neural-mass models (NMMs), this interdependence between mean and variance is dropped, which means that the product of the Hessian $\partial^2 f_l^{(j)} / \partial x^2$ and the covariance matrix $\Sigma^{(j)}$ or the trace of the product has to evaluate to zero (for details, consult Marreiros et al., 2009). While NMMs enjoy much attention because of their relative simplicity, MFMs are capable of representing more complex dynamics (Marreiros et al., 2009). In the following equations, we consider the NMM formulation for cbDCM.

In this case, one simply recovers the deterministic part of Eqs. (7) and (8) (Marreiros et al., 2009):

$$C\dot{\mu}_{V}^{(j)} = \left(\sum_{k} \mu_{g_{k}}^{(j)} \left(V_{k} - \mu_{V}^{(j)}\right)\right) + \mu_{g_{NMDA}}^{(j)} m\left(\mu_{V}^{(j)}\right) \left(V_{NMDA} - \mu_{V}^{(j)}\right)$$

$$k \in L, \text{AMPA, GABA$$
(13)

$$\dot{\mu}_{g_k}^{(j)} = \kappa_k^{(j)} \Big(\zeta_k^{(j)} - \mu_{g_k}^{(j)} \Big)$$
(14)

where the hidden state index *l* from Eq. (11) now takes the values: $l \in \{V, g_k\}$. In addition, $\mu_V^{(j)}$ represents the *j*-th population's mean voltage, whereas $\mu_{g_k}^{(j)}$ is the population mean conductance for channel *k*. We can now express $\zeta_k^{(j)}$, the input to the *j*-th neuronal population, as:

$$\zeta_k^{(j)} = \sum_i \gamma_k^{(j,i)} \sigma\left(\mu_V^{(i)} - V_R, \Sigma^{(i)}\right)$$
(15)

where $\gamma_k^{(j,i)}$ represents a coupling parameter for channel type k, from population i to population j. Moreover, $\sigma()$ is the cumulative distribution function of the univariate normal distribution $\mathcal{N}(\mu_V^{(j)} - V_R, \Sigma^{(j)})$, where V_R represents the threshold potential (-40 mV). Intuitively, $\sigma()$ can be understood as the proportion of active (spiking) afferent neurons. Here, the variance $\Sigma^{(j)}$ is also a free parameter (Moran et al., 2011a).

6. Modeling a source

Now that we are able to model the activity of single neuronal **populations** (as a reminder, we are not modeling single neurons but **ensembles** of neurons), we can define a cortical column (*i.e.*, a cortical functional unit).

The original formulation of DCM for EEG by David et al. (2006a) was based on the model by Jansen and Rit (1995), which was constructed based on previous experimental work with cats and humans (Lopes Da Silva et al., 1974). This initial formulation considered three different neuronal populations: excitatory pyramidal cells, excitatory spiny stellate cells and inhibitory interneurons (Fig. 5). Spiny stellate cells are found in layer IV of the cortical column, whereas the other two cell populations are considered to occupy both the supra- and infragranular layers. Subsequently, a canonical-microcircuit model for DCM was developed (Bastos et al., 2012), which further divides the pyramidal-cell population into two distinct superficial (supra-granular) and deep (infragranular) subgroups, with distinctive connectivity patterns. In this setting, it is therefore possible to differentially model the contribution of each pyramidal cell population to the signal, adding more detail to the model.

The cell populations within a source interact through a set of intrinsic (or within-source) connections, represented in Fig. 5. These connections are defined based on the Jansen-Rit model (Jansen and Rit, 1995), in case of the 3-population model. The intrinsic connections are encoded by the coupling parameters $\gamma_k^{(j,i)}$ from Eq. (15). By fully defining a layered circuit, with a specific set of neuronal populations and intrinsic connections, we have effectively described **one source**.

⁵ Later, we will discuss how to model the so-called neuronal innovations, which represent spontaneous endogenous fluctuations of the neuronal signal at the level of a network of sources.





Fig. 6. Example locations for modeled sources, connected via extrinsic, or between-sources connections. Figure created with Biorender.com.

7. Modeling between-source connectivity

In the previous section, we have explained how one can model a single source. This source is then set to represent a specific brain region: for instance, the medial prefrontal cortex or a particular visual area. However, neural processes typically unfold as the result of interactions in a network of multiple sources. As can be seen in Fig. 6, such a network perspective can be taken by placing sources in several brain areas, which are then connected via weighted extrinsic or between-source connections.

In DCM for electrophysiological data, the definition of the betweensource connections follows a simplified version of the connectivity rules proposed by Felleman and Van Essen (1991). These were derived from experimental studies, most prominently on the monkey visual cortex. **Fig. 5.** Cortical column, based on the Jansen-Rit model (Jansen and Rit, 1995). Spiny stellate cells are found in layer IV of the cortical column, whereas the other two cell populations are considered to occupy both the supra- and infragranular layers. Note that we model populations of cells, not individual neurons (figure modified from Moran et al., 2013). For illustration purposes, the intrinsic glutamatergic connections mediated by the AMPA and NMDA receptors are grouped. However, these can be modeled separately. For the current implementation in SPM12, refer to Section 11. Figure created with Biorender.com.

Extrinsic connections are divided into several types, based on the layers in which they terminate. **Forward connections**, which run from hierarchically lower to higher areas and mainly originate from supragranular layers, terminate in (granular) layer IV, whilst **backward connections**, which run from hierarchically higher to lower brain areas and mostly originate from infragranular layers, terminate in the supra- and infragranular layers, avoiding layer IV. **Lateral connections** terminate in all three layers (see Fig. 7) (David et al., 2006a, 2005).

8. Observation model

To model how the dendritic signal (*i.e.*, activity) from the pyramidal cells gives rise to the modeled signal in the sensors (*e.g.*, EEG electrodes), another layer needs to be defined on top of the neuronal model. This observation model $h(x(t), \theta)$ thus transforms the source signal into the signal that is being modeled, and always includes a forward mapping from sources to sensors. In addition, the particular form of the observation model also depends on the type of signal modeled (*e.g.*, LFP or frequency-domain data).

In its simplest form, the observation model is nothing more than a projection from sources to the sensors. This projection is assumed to be linear and instantaneous and can be expressed as follows (Kiebel et al., 2009):

$$\hat{y}_i(t) = h(x(t), \theta) = L(\theta)x(t)$$
(16)

Here, $\hat{y}_i(t)$ represents the **predicted measurements** at sensor *i*, $h(x(t), \theta)$ the **observation model** with parameters θ , x(t) is the **pyramidal cell activity** and $L(\theta)$ the **lead field or gain matrix**. This matrix describes the passive conduction of the electromagnetic field from sources to sensors and can assume different forms, depending on the source reconstruction technique used as well as on the data modality considered (*i.e.*, EEG, MEG or LFP). Specifically, as discussed in Section 2, two main types of approaches to source reconstruction exist (Daunizeau et al., 2006). The first are called "dipole fit" models, and represent sources with a small number of ECDs (up to five). These models are fast to compute, but provide limited spatial representation of



Fig. 7. Depiction of the several types of extrinsic, or between-source connections. Please note that the vertical arrangement of the 3 types of neurons should not be misinterpreted as statements about layers. While spiny stellate cells are only present in the granular layer (layer IV), pyramidal cells and inhibitory interneurons are located in both supragranular layers (layers II/III) and infragranular layers (layers V/VI). It is assumed that the exogenous input enters a source via the spiny stellate cells in the granular layer (layer IV). In addition, given that the apical dendrites of cortical pyramidal cells are mostly arranged perpendicularly to the cortical surface and in parallel with respect to each other, it is considered that pyramidal cells are the main generators of the measured M/EEG response (Hauk, 2013). For LFP, it is the arrangement and orientation of the dendritic tree versus the soma, along with the parallel arrangement of the apical dendrites, that make pyramidal cells the main generators of signal. As such, in DCM for electrophysiological data, pyramidal cells are assumed to have the most important contribution to the measured signal. Figure created with Biorender.com.

the signal (Daunizeau et al., 2006; Kiebel et al., 2008a). More recently, "imaging" approaches have been formulated, which model sources using a dense set of dipoles. While imaging approaches offer more spatial detail, they also involve estimating a much larger number of parameters, making them more computationally intensive procedures. A detailed discussion of these methods as well as their exact implications for the lead field matrix can be found in earlier publications (Daunizeau et al., 2006; Kiebel et al., 2008b, 2008a, 2006). Furthermore, as mentioned previously, the data modality also influences the form of the lead field matrix. For example, in EEG, it is not uncommon to reduce the dimensionality of the data using, for instance, singular value decomposition. Thus, in this case, the dimensions of the lead field matrix will be: number of components times the number of hidden states. When working with LFP data with few channels that are spatially sufficiently separated, one might forego the dimensionality reduction step and assume that each source is sensed by a single sensor only (*i.e.*, there is no mixing of source signal at the sensor level). In this case, the lead field matrix will be a diagonal matrix with dimensions: number of channels times number of hidden states.

To complete the model of time series data $y_i(t)$, one now only needs to furnish the model with assumptions about the measurement noise $\epsilon \sim \mathcal{N}(0, \Sigma)$:

$$y_i = \hat{y}_i + \epsilon \tag{17}$$

In ERP models of LFP or M/EEG data, predictions and data are dealt with in the time domain. However, one can also summarize the data in the frequency domain by computing what is called the **cross power spectral density** (or **cross-spectral density**, CSD). This allows for a compact representation of long time series measurements and is typically done with "resting-state" data, *i.e.*, a situation without any experimentally controlled stimuli in which brain activity consists of "spontaneous" oscillations. This is the scenario we will consider in the following.

The CSD can be seen as a relationship between two time series as a function of frequency. More specifically, one can think of the CSD as a frequency domain analysis of the covariance between two signals. The CSD $S_{ij}(\omega)$ between two signals $y_i(t)$ and $y_i(t)$ is given by:

$$S_{ij}(\omega) = \mathbb{E}\left[\mathcal{F}\{y_i\}\overline{\mathcal{F}\{y_j\}}\right]$$
(18)

where \mathbb{E} represents the expectation operator, $\mathcal{F}\{y_i\}$ is the Fourier transform of signal $y_i(t)$ and $\overline{\mathcal{F}\{y_i\}}$ the complex conjugate of $\mathcal{F}\{y_i\}$. For the special case i = j, $S_{ij}(\omega)$ is called the **power spectral density** (PSD).

Since electrophysiological measurements are performed using several sensors, we obtain as many time-series as we have channels.⁶ If one computes the CSD for each possible pair of signals $y_i(t)$ and $y_j(t)$, one obtains a symmetric matrix for each considered frequency ω . This means that the full CSD, for all considered frequencies, is a three-dimensional tensor (Fig. 8).

Consequently, if we aim to model data in the frequency domain, we need to add yet another step to the observation model that maps timeseries data (Eq. (17)) to CSD, according to Eq. (18). For a linear system, this is done most efficiently by expressing the mapping between the so-called neuronal fluctuations (or endogenous oscillations; recall that we are in a resting-state scenario) and the measured signal in terms of the kernel of this system (Friston et al., 2012). As we will see in the next section, in this case, the forward mapping simply corresponds to a convolution of the neuronal fluctuations with the derived kernel.

9. Modeling cross-spectral densities

In this section, we will specify a generative model of predicted CSD for resting-state data, based on a mapping between neuronal fluctuations or innovations $u_k(t) \in \mathbb{R}$ (which can be thought of as baseline oscillations of neuronal signal driving the neural populations⁷) and observable signals $y_i(t) \in \mathbb{R}$. Under linear assumptions, this mapping is defined by convolving the neuronal innovations with their corresponding kernel $\kappa_i^k(\tau, \theta)$. A kernel $\kappa_i^k(\tau, \theta)$ is specified as follows (Friston et al., 2012):

$$\kappa_i^k(\tau, \ \theta) = \frac{\partial y_i(t)}{\partial u_k(t-\tau)} \tag{19}$$

⁶ While this is true for the raw (*i.e.*, unprocessed) data, in practice, principal component analysis (PCA) is often used as a preprocessing step to reduce the dimensionality of the data.

 $^{^7}$ Refer to Section 11 for a detailed account of how these fluctuations are modeled in SPM12.



Fig. 8. Expressing time-domain data in the frequency domain. CSD = cross spectral density; PSD = power spectral density. Figure created with Biorender.com.

Here, θ are the parameters we wish to estimate, *k* indicates the *k*-th innovation and *i* the *i*-th channel (Friston et al., 2012). Note that $u_k(t - \tau)$ is defined with a time lag τ . Indeed, there is a delay between the generation of inputs $u_k(t - \tau)$ and that of the signal $y_i(t)$.

Finally, note that, for ease of writing, the Fourier transform will now be indicated by a capital letter (*e.g.*, $K_i^k = \mathcal{F}\{\kappa_i^k\}$). For further simplicity, we follow the notation from Friston et al. (2012) and omit the dependencies on the frequency ω and parameters θ in the following derivations. The predicted CSD are denoted as $S_{ij}(\omega, \theta)$. According to Eq. (18), the predicted CSD $S_{ij}(\omega, \theta)$ is formulated as follows:

$$S_{ij}(\omega, \theta) = \mathbb{E}\left[\mathcal{F}\left\{y_{i}\right\} \cdot \overline{\mathcal{F}\left\{y_{j}\right\}} \right]$$

$$= \mathbb{E}\left[\mathcal{F}\left\{\sum_{k} \kappa_{i}^{k} * u_{k}\right\} \cdot \overline{\mathcal{F}\left\{\sum_{l} \kappa_{j}^{l} * u_{l}\right\}} \right]$$

$$= \mathbb{E}\left[\sum_{k} \sum_{l} K_{i}^{k} \cdot U_{k} \cdot \overline{K_{l}^{l}} \cdot U_{l}\right]$$

$$= \mathbb{E}\left[\sum_{k} \sum_{l} K_{i}^{k} \cdot \overline{K_{l}^{l}} \cdot U_{k} \cdot \overline{U_{l}}\right]$$

$$= \sum_{k} \sum_{l} K_{i}^{k} \cdot \overline{K_{l}^{l}} \cdot \mathbb{E}\left[U_{k} \cdot \overline{U_{l}}\right]$$

$$(20)$$

It is therefore assumed that the CSD can be represented as a sum of convolved neuronal innovations. In the last step, we use the fact that the expectation operator is linear and that U_k and U_l are our random variables of interest (in the frequency domain). In addition, as will be seen later, since the equations of motion are linear in $u_k(t)$, the kernel does not depend on the neuronal innovations, therefore allowing us to pull the first two terms out of the expectation.

The real and imaginary parts of U_k are assumed to be identically and independently distributed (i.i.d.), as follows (Friston et al., 2012):

$$p(Re(U_k)) = \mathcal{N}(0, \gamma_k) \tag{21}$$

$$p(Im(U_k)) = \mathcal{N}(0, \gamma_k) \tag{22}$$

Note that if two random variables *X* and *Y* are independent, then their covariance is zero. This means that: $cov(X, Y) = \mathbb{E}[(X - \mathbb{E}[X])]$ $(Y - \mathbb{E}[Y])] = 0$. Applying this standard definition of the covariance, we obtain the following relation:

$$\operatorname{vov}(\operatorname{Re}(U_k), \operatorname{Re}(U_l)) = \mathbb{E}[\operatorname{Re}(U_k) - \mathbb{E}[\operatorname{Re}(U_k)]] \cdot \mathbb{E}[\operatorname{Re}(U_l) - \mathbb{E}[\operatorname{Re}(U_l)]]$$
$$= \mathbb{E}[\operatorname{Re}(U_k) - 0] \cdot \mathbb{E}[\operatorname{Re}(U_l) - 0]$$
$$= \mathbb{E}[\operatorname{Re}(U_k)] \cdot \mathbb{E}[\operatorname{Re}(U_l)]$$
$$= 0$$
(23)

The same applies to the imaginary component of U_k :

$$\operatorname{cov}(Im(U_k), Im(U_l)) = \mathbb{E}[Im(U_k)] \cdot \mathbb{E}[Im(U_l)]$$

= 0 (24)

In addition, note that, using the linearity of the expectation operator, the expression for $\mathbb{E}[U_k \cdot \overline{U_l}]$ can be rewritten as:

$$\mathbb{E}\left[U_{k} \cdot \overline{U_{l}}\right] = \mathbb{E}\left[\left(Re(U_{k}) + j \cdot Im(U_{k})\right)\left(Re(U_{l}) - j \cdot Im(U_{l})\right)\right]$$
$$= \mathbb{E}\left[Re(U_{k}) \cdot Re(U_{l})\right] - j \cdot \mathbb{E}\left[Re(U_{k}) \cdot Im(U_{l})\right]$$
$$+ j \cdot \mathbb{E}\left[Im(U_{k}) \cdot Re(U_{l})\right] + \mathbb{E}\left[Im(U_{k}) \cdot Im(U_{l})\right]$$
(25)

where *j* denotes the imaginary number: $j = \sqrt{-1}$.

Thus, if $k \neq l$:

$$\mathbb{E}\left[U_k \cdot \overline{U_l}\right] = 0 \tag{26}$$

since the aforementioned i.i.d. assumption also implies that the cross-terms from Eq. (25) are zero. If k = l:

$$\mathbb{E}\left[U_k \cdot \overline{U_k}\right] = \gamma_k + 0 + 0 + \gamma_k$$
$$= 2\gamma_k$$
$$= \lambda_k \tag{27}$$

where we define a new variable λ_k to be the spectral density of the neuronal innovations.

Therefore, making use of Eqs. (26) and (27), Eq. (20) simplifies to:

$$S_{ij}(\omega,\theta) = \sum_{k} \sum_{l} K_{i}^{k} \cdot \overline{K_{j}^{l}} \cdot \mathbb{E} \left[U_{k} \cdot \overline{U}_{l} \right]$$
$$= \sum_{k} K_{i}^{k} \cdot \overline{K_{j}^{k}} \cdot \mathbb{E} \left[U_{k} \cdot \overline{U_{k}} \right]$$
$$= \sum_{k} K_{i}^{k} \cdot \overline{K_{j}^{k}} \cdot \lambda_{k}$$
$$= \sum_{k} S_{ij}^{k}(\omega,\theta)$$
(28)

One hence sees that the predicted cross-spectrum $S_{ij}(\omega, \theta)$ is a linear mixture of the cross-spectra induced by each innovation (Friston et al., 2012). Importantly, rather than making assumptions about the neuronal innovations in the time domain, one can instead directly parameterize the spectral density of the neuronal innovations, λ_k . For this, the following form has been suggested based on previous theoretical and experimental work (Destexhe and Rudolph-Lilith, 2012; Shin and Kim, 2006; Stam and de Bruin, 2004; Stevens, 1972):

$$\lambda_k(\omega) = \alpha + \frac{\beta^{(1)}}{\omega^{\beta^{(2)}}} \tag{29}$$

Here, α represents white noise and the second term colored noise, with $\beta^{(2)}$ indicating the "color" of the noise⁸ and $\beta^{(1)}$ its magnitude. In the most recent version of SPM, this is implemented somewhat differently. We will review precisely how SPM models the spectral density of the neuronal innovations in Section 11 of this paper.

9.1. Specifying the kernel

To be able to generate predictions, we now need to specify the kernels $\kappa_i^k(\tau, \theta)$. These can be computed analytically under further simplifying assumptions. We first define the forward mapping which links endogenous innovations *u* to hidden states *x* and finally to the observed signal y_i (for channel *i*). This mapping contains both the equations of motion $\dot{x} = f(x(t), \theta, u(t))$ and the observation model $h(x(t), \theta)$. For notational simplicity, we omit in the following derivations the dependency of both these functions on θ , *u* and *x*, thus expressing them as $\dot{x}(t)$ and h(t), respectively. Furthermore, we define $J = \partial \dot{x}/\partial x$ as the Jacobian of the neuronal system. Using the chain rule (Friston et al., 2012):

$$\kappa_i^k(\tau, \theta) = \frac{\partial y_i(t)}{\partial u_k(t-\tau)} = \frac{\partial y_i(t)}{\partial h(t)} \cdot \frac{\partial h(t)}{\partial x(t)} \cdot \frac{\partial x(t)}{\partial x(t-\tau)} \cdot \frac{\partial x(t-\tau)}{\partial \dot{x}(t-\tau)} \cdot \frac{\partial \dot{x}(t-\tau)}{\partial u_k(t-\tau)}$$
(30)

By assuming a linear differential equation with Jacobian $J = \partial \dot{x} / \partial x$, we have: $\dot{x}(t) = J \cdot x(t)$

Therefore: $x(t) = \exp(Jt) \cdot C$, where *C* represents the initial conditions of this system. One can hence re-express this equation as: $x(t) = \exp(J\tau) \cdot x(t-\tau)$

Thus, by inserting this last expression into Eq. (30):

$$\begin{aligned} c_i^k(\tau, \ \theta) &= 1 \cdot \frac{\partial h(t)}{\partial x(t)} \cdot \exp(J\tau) \cdot \left(\frac{\partial \dot{x}(t-\tau)}{\partial x(t-\tau)}\right)^{-1} \cdot \frac{\partial \dot{x}(t-\tau)}{\partial u_k(t-\tau)} \\ &= \frac{\partial h(t)}{\partial x(t)} \cdot \exp(J\tau) \cdot \left(\frac{\partial \dot{x}(t)}{\partial x(t)}\right)^{-1} \cdot \frac{\partial \dot{x}(t)}{\partial u_k(t)} \\ &= \frac{\partial h(t)}{\partial x(t)} \cdot \exp(J\tau) \cdot J^{-1} \cdot \frac{\partial \dot{x}(t)}{\partial u_k(t)} \end{aligned}$$
(31)

Note that the omission of the delay τ from the first to the second line is justified by the fact that $\partial \dot{x}(t)/\partial x(t)$ is linearly approximated and that the equations of motion $\dot{x}(t)$ are linear with respect to the inputs. Thus, taking the derivative at one or a later time point is assumed to yield the same result. In addition, because of the second linearity, note that the kernel expression does not depend on $u_k(t)$.

9.2. Adding noise

ŀ

To complete the specification of the forward mapping to crossspectral data, it is presumed that the data is a mixture of the output of the forward mapping (*i.e.*, predicted cross-spectra $S_{ij}(\omega, \theta)$ and channel noise $\lambda^{c}(\omega)$) and Gaussian error $\epsilon_{ij}(\omega)$ (Friston et al., 2012):

$$S_{ij}(\omega) = S_{ij}(\omega, \theta) + \lambda^c(\omega) + \epsilon_{ij}(\omega)$$

= $\sum_k K_i^k \cdot \overline{K_j^k} \cdot \lambda_k + \lambda^c(\omega) + \epsilon_{ij}(\omega)$ (32)

Where it is assumed that:

$$\lambda^{c}(\omega) = \eta + \frac{\nu^{(1)}}{\omega^{\nu(2)}} \tag{33}$$

$$\operatorname{Re}(\epsilon_{ij}) \sim \mathcal{N}(0, \Sigma_{\epsilon})$$
 (34)

$$\operatorname{Im}(\epsilon_{ii}) \sim \mathcal{N}(0, \Sigma_{\epsilon})$$
(35)

The spectral density of the channel noise $\lambda^{c}(\omega)$ is parameterized using a white (η) and a colored component. In addition, the real and imaginary parts of the error term are assumed to follow the same Gaussian distribution, with mean 0 and variance Σ_{e} . Once again, this is implemented somewhat differently in the most recent version of SPM and the exact procedure will be reviewed in the Section 11.

As a final remark, note the conceptual differences between the channel noise and Gaussian error terms: channel noise represents structured noise which is integrated in the forward mapping, and contains parameters which are estimated during model inversion. On the other hand, Gaussian error should be interpreted as measurement noise. This term furnishes the form of the likelihood function (see next section) and is defined by hyperparameters.

10. Model inversion

In the previous sections, we discussed how the forward model of cb-DCM for cross-spectral densities is constructed, *i.e.*, how data could be generated from a set of parameters. Typically, however, we face the inverse challenge: given some measured data *y*, how can one obtain the most probable parameter distribution which could have generated the given data? In the context of generative models, solving this inverse problem is called model inversion or inference. This rests on Bayes' theorem: for a given model *m*, parameter estimation is equivalent to computing the posterior distribution of the parameters θ given the observed data *y* (David et al., 2006a), according to Bayes' rule:

$$p(\theta|y,m) = \frac{p(y|\theta,m)p(\theta|m)}{p(y|m)}$$

where : $p(y|m) = \int p(y|\theta,m)p(\theta|m) d\theta$ (36)

⁸ As can be seen in Eq. (29), in colored noise, the power spectral density is, for positive $\beta^{(2)}$, inversely proportional to the frequency ω of the signal (Szendro et al., 2001).

According to Eq. (36), computing the posterior $p(\theta|y, m)$ requires the likelihood $p(y|\theta, m)$, the prior $p(\theta|m)$ and the normalization constant p(y|m) (also called model evidence). The forward mapping from hidden states to observable signals described in the previous section defines the likelihood function. In addition, priors can be specified over the parameters (for examples, see Marreiros et al. (2010a). However, because the normalization constant (or model evidence) p(y|m) is an integral which can be very difficult to compute, calculating the posterior distribution of the parameters can be seldom done analytically. Instead, model inversion typically relies on approximate inference. Approximate inference approaches are of two types, depending on whether they rely on stochastic or deterministic approximations. Stochastic approaches include Markov Chain Monte Carlo and other sampling techniques, which have the property that they would approach exact results if infinitely many samples could be drawn (Bishop, 2006). Model inversion using sampling methods has been successfully implemented for DCM (Chumbley et al., 2007; Raman et al., 2016; Sengupta et al., 2016, 2015). However, these techniques are associated with increased computational demands. Thus, in the context of DCM, deterministic methods based on parametric approximations to the posterior distribution, despite their potential issues with local optima, are most commonly used. Indeed, the standard inversion technique used in DCM is Variational Bayes (VB). In short, VB approximates the log of p(y|m) (the log model evidence) with a lower bound that is based on a parametric and computable approximation $q(\theta)$ to the posterior. This lower bound is called the negative free energy F(q, y):

$$F(q, y) = \underbrace{\langle \log p(y|\theta, m)_q \rangle}_{\text{Expected log-likelihood}} - \underbrace{D_{KL}[q(\theta) \| p(\theta|m)]}_{\text{Kullback-Leibler divergence}}$$
(37)

The goal is to maximize this quantity F(q, y). However, note how Eq. (37) shows that the negative free energy comprises two distinct terms. Specifically, the expected log-likelihood can be thought of as an accuracy term whereas the Kullback-Leibler (KL) divergence between approximate posterior and prior represents a complexity term (also known as Bayesian surprise) (Friston, 2010). The complexity term effectively serves as a regularization term which acts against overfitting.

VB applied to the inversion of DCMs make two major assumptions. Firstly, it is assumed that $q(\theta)$ can be factorized over disjoint sets of parameters, as follows (Bishop, 2006; Friston et al., 2006):

$$q(\theta) = \prod_{i=1}^{M} q_i(\theta_i)$$
(38)

This is called the mean-field approximation. Under this assumption, it can be shown that the optimal solution $q_i^*(\theta_i)$ can be expressed as (for details, please consult Bishop, 2006; Friston et al., 2006):

$$q_i^*(\theta_i) = \frac{\exp\left(\mathbb{E}_{j \neq i} \left[\log p(y, \theta | m)\right]\right)}{Z^i}$$
(39)

Where $\mathbb{E}_{j\neq i}[\cdot]$ represents the expectation taken over all $q_j(\theta_j)$ where $j \neq i$, and Z^i constitutes a normalization constant. In addition to the mean field approximation, each $q_i(\theta_i)$ is defined to be Gaussian (Laplace approximation) (Friston et al., 2007). We can easily see in Eq. (39) that each $q_i^*(\theta_i)$ is dependent on the other factors $q_{i}^*(\theta_{i})$. Thus, under these approximations, model inversion of DCMs rests on iteratively updating the sufficient statistics of $q_i(\theta_i)$ (*i.e.*, mean and covariance) in order to maximize the negative free energy (Friston et al., 2006; Kiebel et al., 2009). After convergence, F(q, y) represents our best approximation to the log model evidence and, simultaneously, $q(\theta)$ will be the best approximation to the true posterior $p(\theta|y, m)$ (where "best" is to be understood in reference to the chosen form of $q(\theta)$).

Finally, note that CSDs are a special case. Indeed, when dealing with complex-valued data, the negative free energy needs to be expressed in a way that separates the real and imaginary parts, and hence splits the associated prediction errors (Friston et al., 2012). Once this is done,



Fig. 9. SPM GUI. The dropdown menu within the blue box will allow you to choose an observation model, whereas the menu in the red box will permit you to select a neuronal model.

the standard scheme described above can be used to infer the DCM parameters. For more details, please refer to the relevant publication (Friston et al., 2012).

11. Implementation in SPM

SPM is a free and open-source MATLAB software package developed at the Functional Imaging Laboratory (FIL), Wellcome Trust center for Neuroimaging, London. It contains a set of tools to analyze neuroimaging data (e.g., fMRI, EEG, MEG) and can also be used to specify and estimate DCMs. Since the original papers introduced the theoretical underpinnings of the models, various additions and changes have been made in the SPM code. We here provide an overview of some of the relevant aspects of the current implementation of (conductance-based) DCM in SPM. Notably, this section does not attempt to contain an exhaustive list of all changes. Instead, we explore only those aspects that we deem most relevant for the user.

11.1. Choosing your model

If you use the SPM graphical user interface (GUI) and try to define a DCM, you will see a window similar to the one in Fig. 9. You will be

Table 2

SPM options for the **neuronal** model. Verification of which functions are called by which options can be done by consulting the spm_dcm_x_neural.m function.

| Option | Neuronal model selected |
|------------------------------------|---|
| ERP (evoked-response potential) | 3-population convolution-based NMM. |
| SEP (sensory evoked potential) | Variant of the ERP model, with different fixed parameter values (see spm_fx_sep.m). As with the ERP model, this is a 3-population convolution-based NMM. |
| LFP (local field potential) | 3-population convolution-based NMM. This model is conceptually the same as the ERP model, except for the inhibitory within-source connections: the inhibitory population has recurrent |
| | self-connections. These were included to allow for modeling of high-frequency oscillations in the beta band (Moran et al., 2007b). |
| CMC (canonical microcircuit) | 4-population convolution-based NMM |
| NMM (neural mass model) | 3-population conductance-based NMM. |
| MFM (mean field model) | 3-population conductance-based MFM. |
| CMM (canonical microcircuit model) | 4-population conductance-based NMM. |
| NMDA | 3-population conductance-based NMM (defining an MFM is also possible, but not via the GUI). This model includes the NMDA receptor. |
| CMM_NMDA | 4-population conductance-based NMM. This model includes the NMDA receptor. |
| NFM (neural field model) | 3-population convolution-based NFM. Although possible via the GUI, combining this option along with an ERP or CSD observation model does not make sense and will consequently return an error (see spm dcm x neural.m). |

Table 3

SPM options for the **observation** model. Verification of which functions are called by which options can be done by consulting the spm_api_erp.m function.

| Option | Observation model selected |
|---------------------------------|--|
| ERP (evoked-response potential) | Observation model for ERPs. Returns predictions in the time domain. Can be used with all of the above neuronal models, except the NFM. |
| CSD (cross-spectral density) | Observation model for resting-state data. Returns predictions in the frequency domain. Can be used with all of the above neuronal models, except the NFM. |
| TFM (time-frequency model) | DCM of induced cross-spectra. SPM will automatically choose a CMC neuronal model for you (see spm_dcm_tfm.m and spm_fx_cmc_tfm.m). |
| IND (induced responses) | DCM of induced responses. SPM will automatically choose the associated neuronal model for you (see spm_dcm_ind.m and spm_fx_ind.m). This model is not covered in the present publication (for details, refer to Chen et al., 2008). |
| PHA (phase coupling) | DCM for phase coupling. SPM will automatically choose the associated neuronal model for you (see spm_dcm_phase.m and spm_fx_phase.m). This model is not covered in the present publication (for details, refer to Penny et al., 2009). |
| NFM (neural field model) | Can only be used with the NFM neuronal model. |



Fig. 10. Within-source connectivity for the 3-population source model, as implemented in the spm_fx_mfm.m and spm_fx_nmda.m functions. Figure created with Biorender.com.

confronted with many options for both the observation and the neuronal models. To facilitate usage and provide guidance, we have listed the available options in SPM12 in Tables 2 and 3 and explain their meaning with regard to the framework introduced in Section 2 of this paper.

11.2. Modeling intrinsic connectivity

The intrinsic connectivity rules for the conductance-based models implemented in the current version of SPM differ from what was initially described by David et al. (2006a). Figs. 10 and 11 illustrate the

Fig. 11. Within-source connectivity for the 4-population source model, as implemented in the spm_fx_cmm.m and the spm_fx_cmm_NMDA.m functions. In the latter function, the inhibitory connection from the superficial pyramidal cell population to the spiny stellate cells is not represented (this is indicated by a less strong line in the figure). Figure created with Biorender.com.

within-source connectivity patterns currently implemented in SPM, with reference to the relevant functions.

11.3. Modeling the neuronal innovations

Within SPM, the spectral density of the neuronal innovations λ_k is modeled somewhat differently from what was discussed previously (Eq. (29)). In SPM, 2 parameter classes are used to this effect: **a** and **d** (see spm_ssr_priors.m and spm_csd_mtf_gu.m). The former models $1/\omega$



Fig. 12. SPM GUI. This window appears during model inversion. The upper plots visually contrast data (dotted lines) and model predictions (solid lines). Note that real and imaginary parts of the data are plotted separately. On the lower panel, the graph displays the difference between conditional (*i.e.*, posterior) and prior expectation, as a function of the parameter (x axis). Notice how some parameters' posterior estimates distance themselves from the prior values, while others do not.

(colored) noise, whereas the latter contains the coefficients for a set of cosine functions. These cosine functions form a basis set that allow for flexibly estimating structure in the neuronal innovations that deviates from the $1/\omega$ component, while keeping the representation of this structure compact (compare the common use of discrete cosine transform basis functions in image processing). This basis function approach to modeling endogenous fluctuations is, to our knowledge, not mentioned explicitly in the published literature on DCM for cross-spectral data; however, it is implemented in the more recent versions of SPM.

In SPM, a is defined as a matrix:

$$\mathbf{a} = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1n} \\ a_{21} & a_{22} & \cdots & a_{2n} \end{bmatrix}$$
(40)

where *n* represents the number of sources in the DCM. As for the first dimension of **a**, for each source *i*, the colored noise $(1/\omega)$ component $\lambda_a^{(i)}(\omega)$ is described as:

$$\lambda_a^{(i)}(\omega) = \exp(a_{1i}) \cdot \omega^{-\exp(a_{2i})} \tag{41}$$

 $\lambda_a^{(i)}(\omega)$ is then a scalar value. If one does this operation for all sources, one obtains the 1 × *n* vector $\lambda_a(\omega)$ for each frequency ω :



Fig. 13. Models used in the simulations. On the left, model 1 includes one single source. On the right, model 2 contains 2 sources. These sources are coupled via one single extrinsic forward connection, from source 1 to source 2. Figure created with Biorender.com.

$$\lambda_{a}(\omega) = \begin{bmatrix} \lambda_{a}^{(1)}(\omega) \\ \lambda_{a}^{(2)}(\omega) \\ \vdots \\ \lambda_{a}^{(n)}(\omega) \end{bmatrix}^{T}$$
(42)

d is also a matrix, of the following form:

$$\mathbf{d} = \begin{bmatrix} d_{11} & d_{12} & \cdots & d_{1n} \\ d_{21} & \ddots & \ddots & d_{2n} \\ \vdots & \ddots & \ddots & \vdots \\ d_{41} & d_{42} & \cdots & d_{4n} \end{bmatrix}$$
(43)

Note that the first dimension indicates the number of cosine functions used to model λ_k and that it can be arbitrarily defined. In the current version of SPM, this number has been set to 4, while previous versions worked with 8 cosine functions.¹⁰ Once again, the parameters are defined in a source-specific manner. The component $\lambda_d(\omega)$ of the spectral density of the neuronal innovations is modeled for all sources using matrix multiplication.

$$\lambda_d(\omega) = \exp(C(\omega) \times d) \tag{44}$$

Here, $C(\omega)$ is a row vector containing the contribution (weight) of each cosine function for frequency ω . In newer versions of SPM with 4 cosine functions, this is therefore a 1 × 4 vector. In addition, exp() indicates the elementwise exponential. Therefore, $\lambda_d(\omega)$ also constitutes a 1 × *n* vector for each frequency ω .

The $1/\omega$ and cosine components are then combined using an elementwise product:

$$\lambda(\omega) = \lambda_a(\omega) \circ \lambda_d(\omega) \tag{45}$$

where $\lambda(\omega)$ corresponds to the vector of neuronal innovations indexed by *k* in Eq. (32).

11.4. Modeling noise

We now address the channel noise terms introduced in Eqs. (32) and (33). In SPM, the colored channel noise added to the predicted CSD is subdivided into what is called non-specific channel noise (which describes the contribution of common noise sources, *e.g.*, in a common reference channel) and specific channel noise (Moran et al., 2009). The white noise component η (Eq. (33)) is set to zero; yet, white noise can, in principle, still be represented, as we will see below. Mathematically, the model is expressed as follows:

$$S_{ij}(\omega) = \sum_{k} K_i^k \cdot \overline{K_j^k} \cdot \lambda_k + S_{ij}^{noise}(\omega) + \epsilon_{ij}(\omega)$$
(46)

¹⁰ Note that these cosine functions are modeling frequency-domain (and not time-series) data. Thus, these frequencies cannot be equated to the frequencies of the time-domain data. Furthermore, note that using a linear combination of cosine functions is a data-driven process. Thus, the range of frequencies picked up depend on the characteristics of the data.

Table 4

Priors distributions used for the simulations: the parameters are defined using a log normal distribution, in order to satisfy certain positivity and negativity constraints. For convenience, we provide the names of the functions where the priors are defined, as well as the functions where these parameters play a relevant role and are used. Abbreviations: ss = spiny stellate cell population, sp = superficial pyramidal, ii = inhibitory interneuron and dp = deep pyramidal cell populations.

| Parameter ϕ_i $\phi_i = \pi_i \exp(\theta_i)$ $\theta_i = \mathcal{N}(0, \sigma_i^2)$ | Meaning and relevant equation | Prior mean | Prior variance σ_i^2 | Relevant SPM functions |
|--|---|---|---|--|
| $\Sigma^{(j)}$ | Variance of univariate normal distribution $\mathcal{N}(\mu_V^{(j)} - V_R, \Sigma^{(j)})$ (Eq. (15)) | $\pi_{Sigma} = 32$ | 1/64 | spm_cmm_NMDA_priors.m spm_fx_cmm_NMDA.m |
| κ | Ion channel rate constant (Eq. (8)) | $\pi_{AMPA} = 1000/4$ $\pi_{GABA} = 1000/16$ $\pi_{NNDA} = 1000/100$ | 1/64 | spm_cmm_NMDA_priors.m spm_fx_cmm_NMDA.m |
| С | Membrane capacitance (Eq. (5)) | $\pi_{ss} = 128/1000$ $\pi_{sp} = 128/1000$ $\pi_{ii} = 256/1000$ $\pi_{dn} = 32/1000$ | 1/16 | spm_cmm_NMDA_priors.m spm_fx_cmm_NMDA.m |
| $\gamma_k^{(j,i)}$ | Coupling parameter for within-source connectivity (Eq. (15)) | $\pi_{\gamma} = 0$ to 128, depending on the exact connection | 1/32 (if connection defined by modeler) or 0 (if connection is not defined) | spm_cmm_NMDA_priors.m spm_fx_cmm_NMDA.m |
| <i>a</i> ₁ | Amplitude of the colored component of the neuronal innovations (Eq. (41)) | $\pi_{a_1} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |
| <i>a</i> ₂ | Exponent of the colored component of the neuronal innovations (Eq. (41)) | $\pi_{a_2} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |
| d | Weight of cosine function (Eq. (44)) | $\pi_d = 1$ | 1/128 | spm_ssr_priors.m spm csd mtf gu.m |
| <i>b</i> ₁ | Amplitude of non-specific channel noise (Eq. (48)) | $\pi_{b_1} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |
| b_2 | Exponent of non-specific channel noise (Eq. (48)) | $\pi_{b_2} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |
| <i>c</i> ₁ | Amplitude of specific channel noise (Eq. (48)) | $\pi_{c_1} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |
| <i>c</i> ₂ | Exponent of specific channel noise (Eq. (48)) | $\pi_{c_2} = 1$ | 1/128 | spm_ssr_priors.m spm_csd_mtf_gu.m |

Here, we have substituted $\lambda^{c}(\omega)$ with $S_{ij}^{noise}(\omega)$. In SPM, two parameter classes model channel noise: **b** and **c**. Each of these parameter classes is a 2 × 1 vector:

$$\mathbf{b} = \begin{bmatrix} b_1 \\ b_2 \end{bmatrix} \quad \mathbf{c} = \begin{bmatrix} c_1 \\ c_2 \end{bmatrix} \tag{47}$$

where **b** contains the parameters that model non-specific channel noise, whereas **c** models specific channel noise. Non-specific noise affects the full cross-spectra, whereas specific noise only affects the power spectra. $S_{ii}^{noise}(\omega)$ is then expressed as:

$$S_{ij}^{noise}(\omega) = \begin{cases} \exp(b_1 - 2) \cdot \omega^{-\exp(b_2)} + \exp(c_1 - 2) \cdot \omega^{-\exp(c_2)} & \text{if } : i = j \\ \exp(b_1 - 2) \cdot \omega^{-\exp(b_2)} & \text{otherwise} \end{cases}$$

$$\tag{48}$$

Note that **b** and **c** are not single-channel-specific and that, in the extreme case where b_2 or c_2 take very small negative numbers, both $-\exp(b_2)$ and $-\exp(c_2)$ will be close to zero and $S_{ij}^{noise}(\omega)$ will essentially model white noise. The exponential functions in the expression ensure positivity and negativity constraints in the magnitude and exponent, respectively, of the noise terms.

11.5. Inverting models

Inverting DCMs can be done in SPM either via the GUI or programmatically. While using the GUI might at first be a more intuitive approach, inverting models at scale is best done by writing a script that automatically fits DCMs over, for instance, multiple subjects, or different conditions. To showcase model inversion, we have used 2-channel cross-spectral LFP data from an anaesthetized rodent. This dataset is freely available and can be downloaded from the SPM website: https://www.fil.ion.ucl.ac.uk/spm/data/dcm_csd. In addition, the SPM developers have written a tutorial on model inversion using this dataset and the GUI (for details, see Chapter 45 of the SPM12 Manual, Ashburner et al. 2018). For this publication, we have written a short script to invert a simple 2-region DCM on these empirical data. The code is freely available and can be found at: https://gitlab.ethz.ch/tnu/code/pereiraetal_conductance_based_dcm. In SPM, DCMs are specified using MATLAB structure arrays. This is the construct one uses to tell SPM which neuronal model and which observation model are to be used, as well as to specify how many sources one wishes to define and their associated connectivity. Once this is done, it is possible to call spm_dcm_csd.m, a "master function" which estimates the posterior parameter estimates of the model. The actual model inversion is performed by the spm_nlsi_GN.m function, which fits the generative model using the procedure described in Section 10. Once the inversion is started, SPM will open a new window

In addition, during model inversion, one sees an output similar to the following in the MATLAB command window:

similar to the one shown in Fig. 12:

| Inverting DCM model | | | | | | |
|-----------------------|---------------|-----------------|-------------|-----------|-----------|-------|
| Using existing priors | | | | | | |
| | | | | | | |
| evaluating CSD | for condition | 1 | | | | |
| EM:(+): 1 sec) | F: 0.000e+00 | dF predicted: | 1.308e+05 | actual: | 6.206e+04 | (3.80 |
| EM:(+): 2 sec) | F: 6.206e+04 | dF predicted: | 1.610e+03 | actual: | 2.368e+03 | (3.60 |
| EM:(+): 3 sec) | F: 6.443e+04 | dF predicted: | 1.501e+03 | actual: | 1.738e+03 | (4.23 |
| [] | | | | | | |
| EM:(+): 45 | F: 6.832e+04 | dF predicted: 1 | .595e-02 co | onvergenc | e | |



Fig. 14. Simulated data using model 1, as specified in Fig. 13. The prior parameter values from SPM12 were used to perform the simulations of plots A, B and C. For these three plots, the value of each parameter was altered up to two prior standard deviations away from the prior mean. $+2\sigma$ indicates that two standard deviations were added to the prior parameter value, whereas -2σ signals that two standard deviations were removed. In addition, data were also simulated using the prior value for the parameter. Only the magnitude of the PSD is plotted. (A) Changes in the value of κ_{AMPA} are more strongly reflected in the 15–30 Hz interval. (B) Increases in the weight of the connection from the spiny stellate cells to the superficial pyramidal cells gives rise to a predicted increase in the magnitude of the spectra, which is more evident below 20 Hz. (C) Changes in the value of parameter $\Sigma^{(f)}$ lead to alterations similar to those observed in subplot A. (D) This subplot is somewhat different from the other three. Here, the magnesium switch was parameterized and α_{NMDA} was allowed to vary. However, since no prior distribution is defined within SPM12, the values tested were manually chosen (see legend), and fall within those evaluated in a previous publication (Moran et al., 2011a). The white dotted line corresponds to the default value in SPM12. α_{NMDA} plays an important role by determining which frequency band is associated with the maximal magnitude.

Note how the command window indicates the number of the iteration, whether the value of the negative free energy has increased (denoted by "(+)") or decreased ("(-)"), the current value for the negative free energy, as well as the predicted and actual change in the negative free energy value at each iteration. When the difference between negative free energy values at successive iterations falls below a determined threshold (in SPM, this threshold is 0.01), the algorithm is thought to have converged and the DCM has been fitted. All results are stored within the DCM Matlab structure array and one can now inspect DCM.Ep for the expectations of the posterior parameter estimates, as well as DCM.Cp for the posterior covariance matrix.

12. Simulations

So far, we have reviewed the mathematical equations for cbDCM for cross-spectral densities derived from electrophysiological data. In addition, we discussed some aspects of the implementation in SPM that we deemed useful for the reader. In this section, we conclude this second, more practical part of the paper by making use of simulations to provide an intuitive and qualitative understanding of the role certain parameters play in the generation of cross-spectra.

To do this, we define two simple models (Fig. 13) of cross-spectral LFP data. Both models are cbDCMs with four neural populations per source and the NMDA receptor. The first model contains only one source to illustrate the effect of within-source parameters. The second model includes two sources to illustrate the effect of between-source parame



Fig. 15. Simulated data, now changing the value of a_2 , the exponent for the colored noise component of the neuronal innovations.

ters. Specifically, in the second model, one single extrinsic connection is defined, from source 1 to source 2.

To examine the effect of changing specific parameter values on the CSD, we start by taking the prior mean, as defined in SPM12, to generate data. We then change the value of single parameters using the standard deviation of the prior distribution of that parameter (*i.e.*, the prior standard deviation). Specifically, we alter the value of single parameters up to two prior standard deviations away from the prior mean. Note that parameters in SPM12 are generally defined in log normal space and subsequently scaled. Thus, for each parameter $\phi_i: \phi_i = \pi_i \exp(\theta_i)$, where $\theta_i = \mathcal{N}(0, \sigma_i^2)$. The scaling factor π_i is responsible for transforming the parameter into the actual physical quantity of interest. Table 4 reproduces the sufficient statistics of the prior distributions used in SPM12 and for these simulations.

In one set of simulations, we changed the current parameterization in SPM12 to illustrate the effect of the magnesium nonlinearity. While the magnesium block parameters are fixed by default in the current SPM implementation, they have been parameters of interest in several previous studies. For example, Moran et al. (2011a) parameterized the magnesium switch as in Eq. (6). In this publication, we also define α_{NMDA} as a parameter of interest and, instead of defining a prior distribution and using the method outlined above, we set it to values similar to those previously tested in this publication (Moran et al., 2011a).

In the following, we plot the synthetic data. Please note that the choice of frequency interval, as well as the parameters that are tested and the range over which the parameter values are varied represents an exemplary selection. Furthermore, we here deliberately defined simple models by restricting ourselves to models with 1 to 2 sources and LFP data. Our aim is to provide an illustration of the dynamics of the models and not to comprehensively test the model or explore what cbDCM can or cannot do. The complexity of cbDCM means that it is not always easy to predict how changes in one parameter will affect the spectra. Thus, if the reader wishes to inspect other parameter settings, the code for these simulations is freely available and can be accessed here: https://gitlab.ethz.ch/tnu/code/preriaretal_conductance_based_dcm. This is meant to offer an easy, hands-on experience with cbDCMs and we encourage the reader to experiment with the code.

12.1. Single-source model

We start by analyzing the parameters of the neuronal model. Fig. 14 shows the results of the simulations when one changes the values of four different parameters (please note that the *y*-axes of the plots are scaled differently). In the context of this first single-source LFP model, we assume only one electrode is present, and therefore have only one channel (Fig. 8), and only one plot per parameter tested.

Let us start with $\kappa_{AMPA},$ the AMPA receptor rate constant (Eq. (8), subplot A in Fig. 14). Visually, one sees that changes in this parameter's value lead to more apparent alterations in the peak around 20 Hz, with decreases in the value of κ_{AMPA} leading to an increase in the magnitude of the PSD. Subplot B is similar to the previous plot, with the difference that the parameter changed is the within-source coupling parameter $\gamma_{\iota}^{(j,i)}$ (Eq. (15)), weighting the connection from the spiny stellate cells (i) to the superficial pyramidal cells (*j*). This particular connection was chosen because the superficial pyramidal cells constitute the neural population that is assumed (within SPM12 for 4-population conductancebased models) to most significantly contribute to the measured signal (for a detailed discussion on the contribution of the different populations, refer to Murakami et al. 2006). Increases in the strength of this connection would intuitively lead to an increase in the magnitude of the spectra, which is what is observed in Fig. 14. This effect is much more pronounced for frequencies below 20 Hz.

Parameter $\Sigma^{(j)}$ (the variance of the univariate normal distribution from Eq. (15)) is assumed to take on the same value for all neuronal populations in the model. When testing this parameter, we obtained an effect qualitatively similar to that observed for κ_{AMPA} (subplot C). Finally, we turn to α_{NMDA} , the magnesium non-linearity parameter, which exerts a profound influence on the PSD, and determines which frequency band is associated with the maximal magnitude. In the current implementation of SPM12, α_{NMDA} is fixed to 0.06, yet has been used as a free parameter in various previous publications (Moran et al., 2011a; Symmonds et al., 2018). In our simulations (subplot D), values below 0.06 are associated with maximal PSD in the lower frequencies (< 5 Hz). Conversely, as the value of α_{NMDA} is increased, a significant rise in the value of the spectra in the 15 to 25 Hz interval occurs, with the beta band (13–29 Hz) eventually becoming the band associated with the maximal PSD.

Fig. 15 displays the data generated by modifying the value of a_2 , the exponent of the colored noise component of the neuronal innovations (Eq. (41)). Changes in a_2 lead to a qualitative effect again similar to that seen for κ_{AMPA} and $\Sigma^{(j)}$, albeit with more pronounced alterations in the frequencies below 20 Hz.

Changes in the cosine set parameters of the neuronal innovations also lead to interesting alterations. For this set of simulations, we change the value of a single parameter up to 5 prior standard deviations away from the prior mean, to make the qualitative effect more apparent. Fig. 16 shows how changes in the value of the 4th cosine function coefficient lead to differences in the strength of an oscillation on top of the prior spectral density. In particular, with more extreme values of the parameter, one can observe more pronounced oscillations around the prior value (0 σ).

Finally, we investigated the effect of the channel noise. Fig. 17 shows the data simulated after changing the value of c_2 , the exponent of the specific, colored channel noise (Eq. (48)). It is evident that under our current settings, the changes induced in the power-spectra are subtle. By inspecting the full band, one does not see any changes. Even after "zooming in" on very high frequencies, the effect still appears very modest. This is likely due to the fact that, under the prior settings, the amplitude of the channel noise is given by exp(-2) (compare Eq. (48)).

12.2. Two-source model

The two-source model was defined specifically to study the effect of the extrinsic connectivity parameters on the power- and cross-spectral densities. As pointed out previously, this second model incorporates 2 sources. These are connected via one single forward projection from source 1 to source 2. Again, by modeling LFPs, we can assume one electrode per source, therefore having a total of two channels (Fig. 8). In



Fig. 16. Same as Fig. 15, now changing the value of d_{41} , the 4-th cosine function coefficient for source 1. The left plot shows the results over the full band, whereas the right plot "zooms in" on the lower frequencies (1–13 Hz), for a clearer visualization of the effect the parameter has on model output.

Fig. 17. Same as Fig. 15, now changing the value of c_2 , the exponent of the specific colored channel noise.

Fig. 18, we simulate data after changing the value of the single AMPAmediated extrinsic connectivity parameter. Only the upper half of the panels is shown, given the symmetry in the data. The magnitude of the complex spectra is plotted and the panel titles indicate which channels are being considered. As expected, the extrinsic connectivity parameter influences the scaling of the spectra, with higher values being associated with increased magnitude of the CSD. Note that the cross-spectra in panel CSD 1/1 are not affected by changes in this parameter. This is because we define no backward connection from Source 2 (Channel 2) to Source 1 (Channel 1). Therefore, changes in the connection from source 1 to Source 2 will only elicit changes in the signal from Source 2, and will not affect the signal from Source 1.

To summarize, while all of the parameters are associated with changes in the model output, the qualitative and quantitative degree





Fig. 18. Simulated data using model 2, as specified in Fig. 13. The prior parameter values from SPM12 were once again used to perform the simulations. In this setting, the weight of the AMPA-mediated extrinsic connection from source 1 to source 2 was altered, up to two prior standard deviations away from the prior mean. Only the upper half of the panels is shown, given the symmetry in the data. The panel titles indicate which channels are being considered. Only the magnitude of the cross-spectra is plotted. Figure created with Biorender.com.

of change vary depending on the parameter considered. For instance, while c_2 exerted a very modest effect on the cross-spectra, other parameters, such as a_2 , had a much more marked effect on the model output. In addition, certain parameters seem to preferably affect one frequency band over another. For instance, while the intrinsic connectivity parameter affected mostly the lower frequencies, other parameters, such as κ_{AMPA} , exerted a greater influence on the beta band. Finally, certain parameters showed an interesting qualitative effect, such as the cosine component of the neuronal innovations and, in particular, the magnesium nonlinearity parameter α_{NMDA} .

13. Summary and conclusions

With the increasing use of more complex and sophisticated models in neuroscience, understanding the conceptual and mathematical principles behind these models has become more challenging. DCM for electrophysiological data, which was the focus of the present paper, builds on previous neurophysiological models and works with several layers of approximation. Keeping track of these steps is a difficult but crucial task if one is to understand the foundations of this class of generative models and be able to judge their strengths but also limitations.

In this paper, we have discussed the different DCM variants for electrophysiological data and reviewed cbDCM in particular. We have described the neuronal model in detail, explaining how single-neuron dynamics can be incorporated in a microcircuit model. We then turned to the observation model and derived it for resting-state electrophysiological data (CSDs).

In a second, more practical part of this paper, we reviewed several aspects related to the current implementation of these models in SPM. In particular, we discussed how intrinsic connectivity is represented for the 3- and 4-population neuronal models, and detailed how neuronal

innovations and channel noise are modeled. Finally, we presented simulations from very simplistic models, in order to equip the reader with a first qualitative understanding of how changes in specific parameters can alter the model's output.

We hope this tutorial paper will prove useful not only for readers starting to work with DCM for electrophysiological data, but also to scientists with more experience with these models.

Declaration of Competing Interest

None

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