

Classical (frequentist) inference

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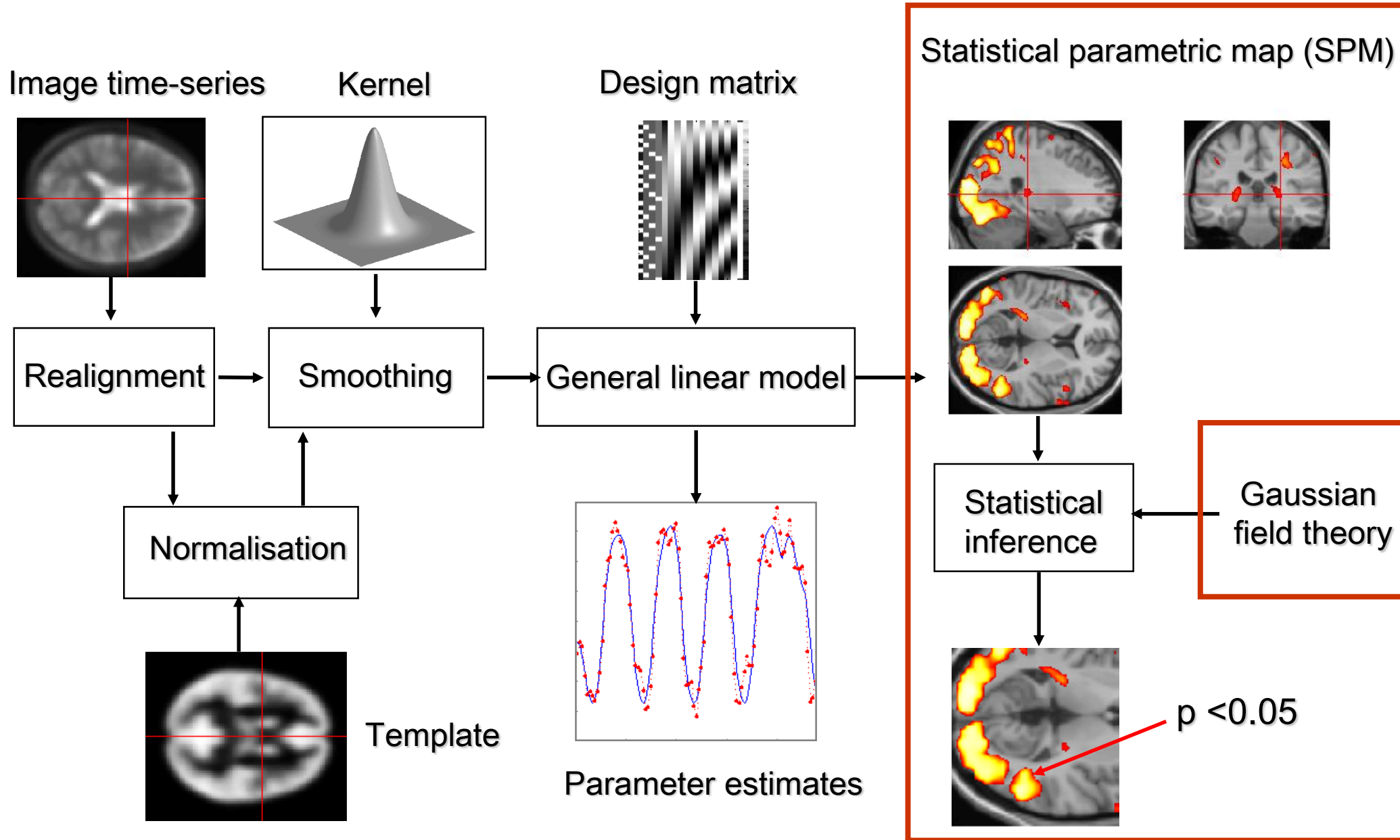


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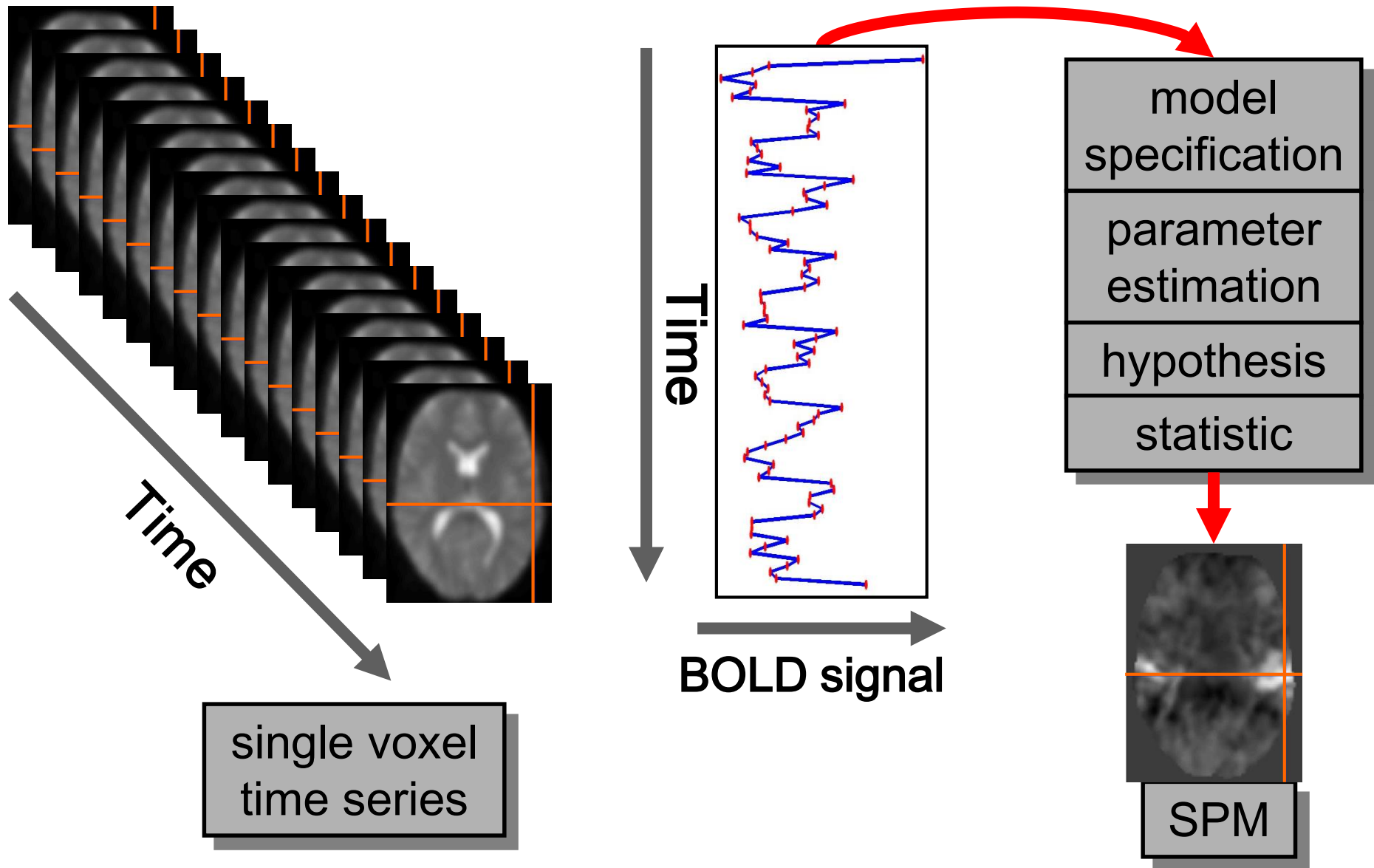
Methods & models for fMRI data analysis
15 October 2019

With many thanks for slides & images to: FIL Methods group

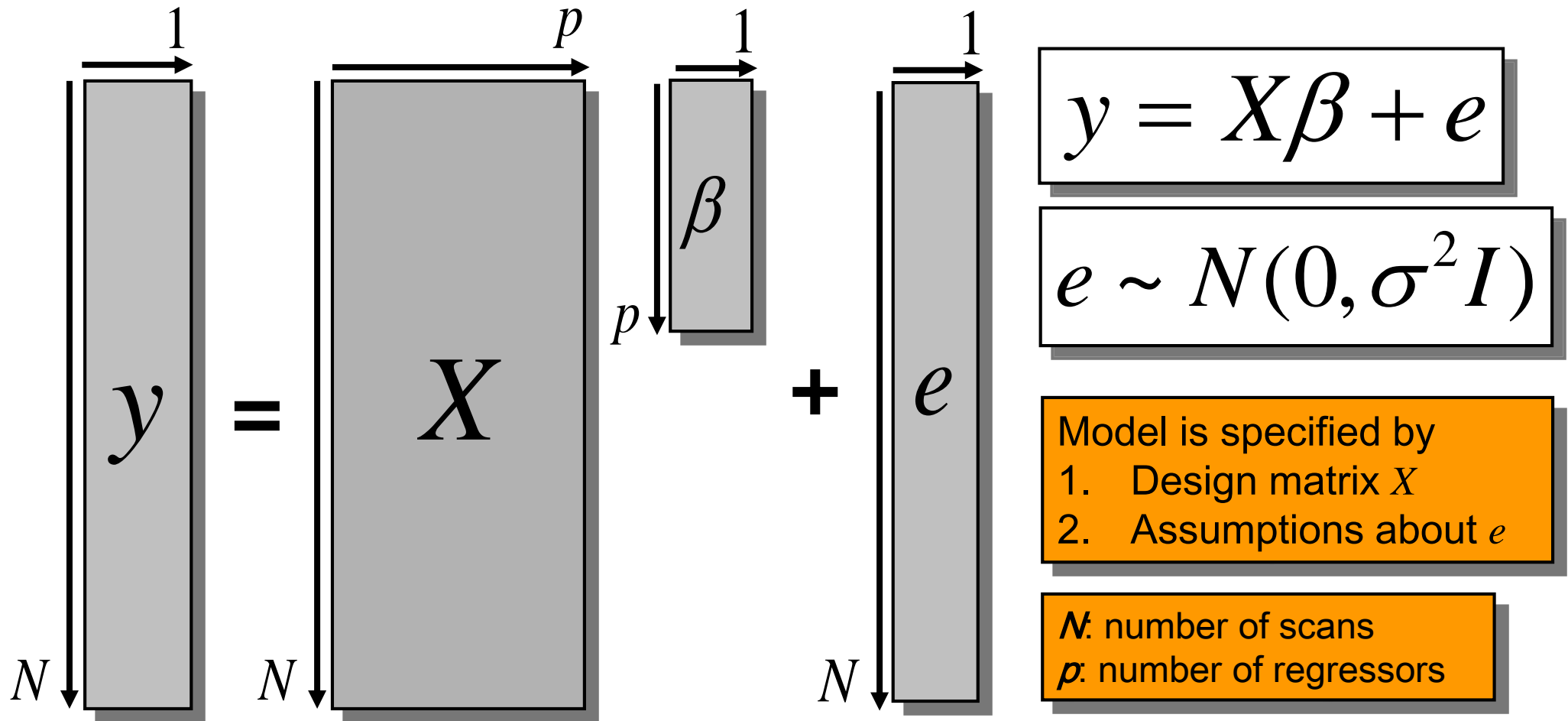
Overview of SPM



Voxel-wise time series analysis



Mass-univariate analysis: voxel-wise GLM



The design matrix embodies all available knowledge about experimentally controlled factors and potential confounds.

Ordinary least squares (OLS) parameter estimation

$$y = X\beta + e$$

The diagram shows a vertical vector y on the left, followed by an equals sign, then a vertical matrix X with alternating black and white horizontal bars. To the right of X is a column vector $\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix}$, followed by a plus sign and another vertical vector e with gray horizontal bars. Below each of these elements are their respective labels: y , X , and e .

$$y = X\beta + e$$

Objective:
estimate parameters
to minimize

$$\sum_{t=1}^N e_t^2$$

Ordinary least squares
estimation (OLS)
(assuming i.i.d. error):

$$\hat{\beta} = (X^T X)^{-1} X^T y$$

Terminology

- A **statistic** is the result of applying a mathematical function to a **sample** (set of data).
- (More formally, a **statistic** is a function of a sample where the function itself is independent of the sample's distribution. The term is used both for the function and for the value of the function on a given sample.)
- A statistic is distinct from an unknown statistical **parameter**, which is a population property and can only be estimated approximately from a sample.
- A statistic used to estimate a parameter is called an **estimator**.
For example, the sample mean is a statistic and an estimator for the population mean, which is a parameter.

Hypothesis testing

To test an hypothesis, we construct a “test statistic”.

- **“Null hypothesis”** H_0 = **“there is no effect”** $\Rightarrow c^T \beta = 0$

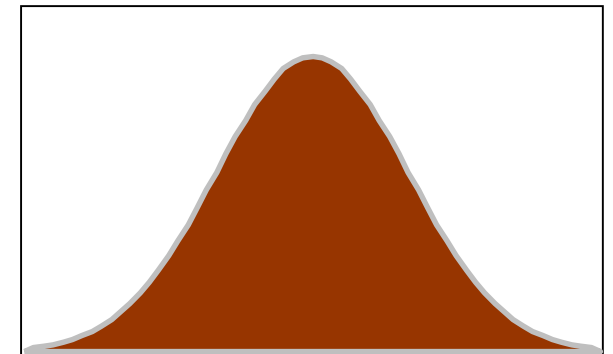
This is what we want to disprove.

\Rightarrow The “alternative hypothesis” H_1 represents the outcome of interest.

- **The test statistic T**

The test statistic summarises the evidence for H_0 .

\Rightarrow We need to know the distribution of T under the null hypothesis.



Null Distribution of T

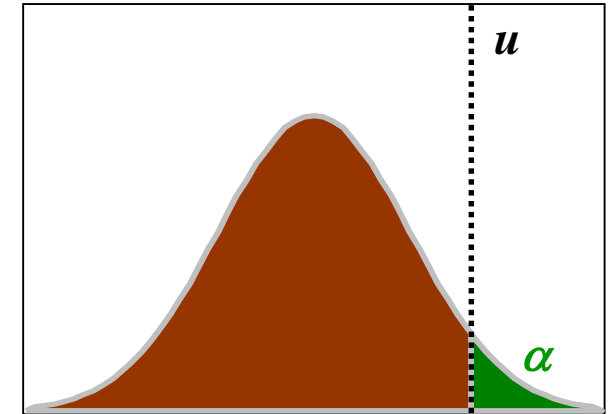
Hypothesis testing

- **Type I Error α :**

Acceptable *false positive rate* α .

Threshold u controls the false positive rate

$$\alpha = p(T > u | H_0)$$



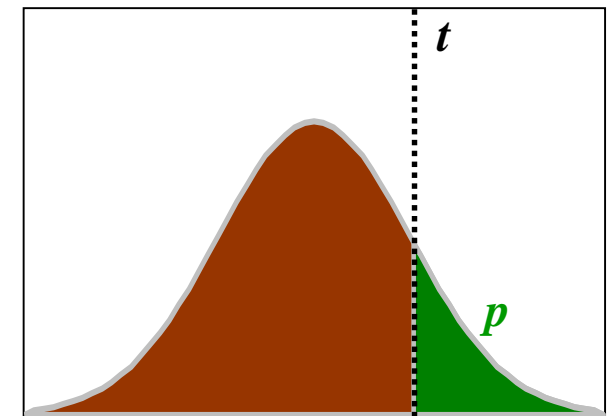
Null Distribution of T

- **Observation of test statistic t , a realisation of T :**

A p -value summarises evidence against H_0 .

This is the probability of observing t , or a more extreme value, under the null hypothesis:

$$p(T \geq t | H_0)$$



Null Distribution of T

- **The conclusion about the hypothesis:**

We reject H_0 in favour of H_1 if $t > u$

Types of error

Actual condition			
		H_0 true	H_0 false
Test result	Reject H_0	False positive (FP) Type I error α	True positive (TP)
	Failure to reject H_0	True negative (TN)	False negative (FN) Type II error β
		specificity: $1-\alpha$ = $TN / (TN + FP)$ = proportion of actual negatives which are correctly identified	sensitivity (power): $1-\beta$ = $TP / (TP + FN)$ = proportion of actual positives which are correctly identified

**One cannot accept the null hypothesis
(one can only fail to reject it)**



Absence of evidence is not evidence of absence!

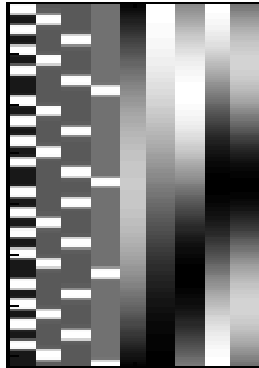
If we do not reject H_0 , then all we can say is that there is not enough evidence in the data to reject H_0 . This does not mean that we can accept H_0 .

What does this mean for neuroimaging results based on classical statistics?

A failure to find an “activation” in a particular area does not mean we can conclude that this area is not involved in the process of interest.

Contrasts

- We are usually not interested in the whole β vector.
- A contrast $c^T\beta$ selects a specific effect of interest:
 - \Rightarrow a contrast vector c is a vector of length p
 - $\Rightarrow c^T\beta$ is a linear combination of regression coefficients β



$$c^T = [1 \ 0 \ 0 \ 0 \ 0 \ \dots]$$

$$c^T\beta = 1\beta_1 + 0\beta_2 + 0\beta_3 + 0\beta_4 + 0\beta_5 + \dots$$

$$c^T = [0 \ -1 \ 1 \ 0 \ 0 \ \dots]$$

$$c^T\beta = 0\beta_1 + -1\beta_2 + 1\beta_3 + 0\beta_4 + 0\beta_5 + \dots$$

- Under i.i.d assumptions:

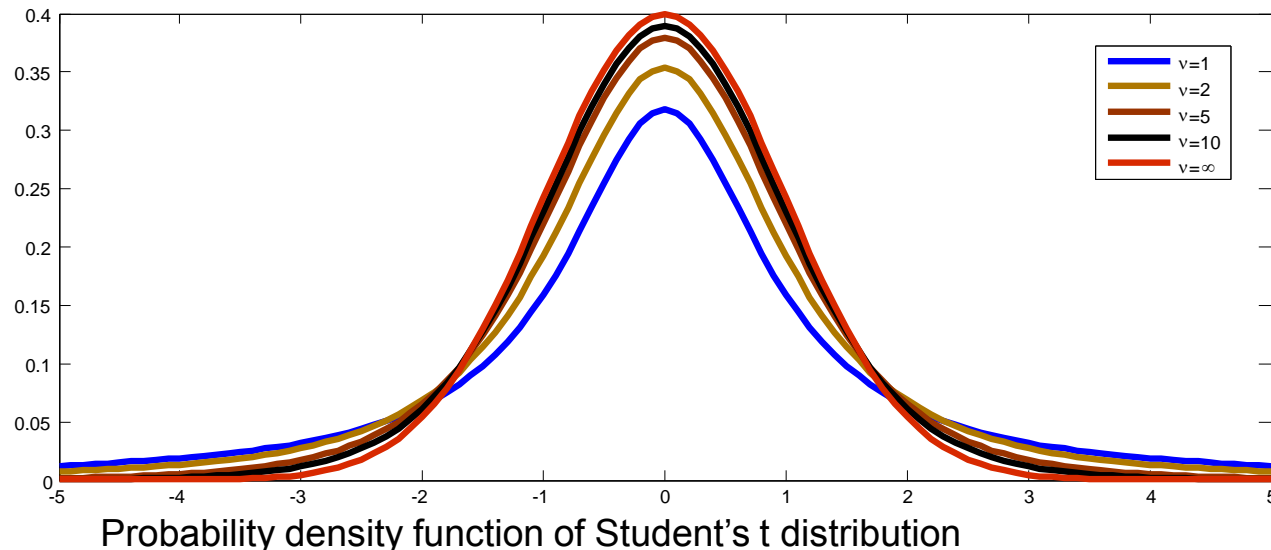
$$c^T \hat{\beta} \sim N(c^T \beta, \sigma^2 c^T (X^T X)^{-1} c)$$

NB: the precision of our estimates depends on design matrix and the chosen contrast !

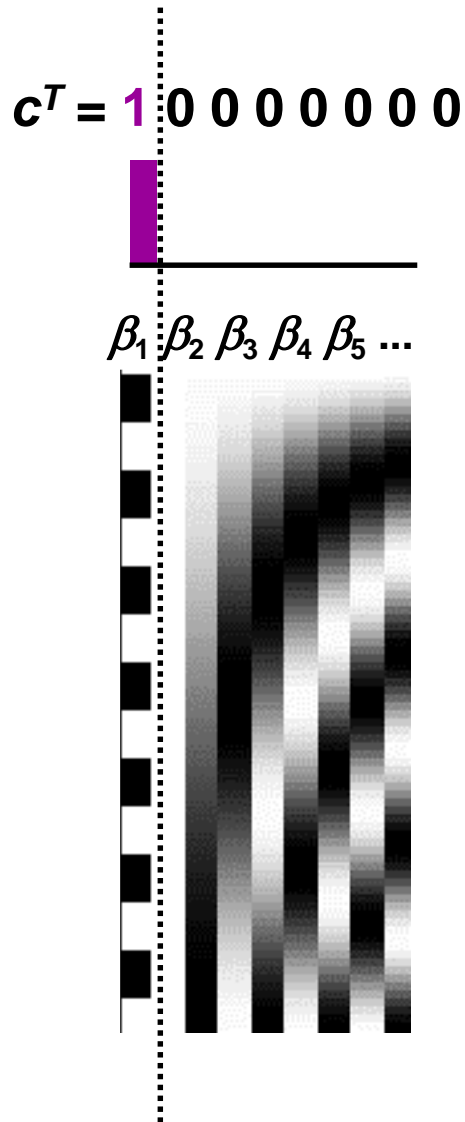
Student's t-distribution

- first described by William Sealy Gosset, a statistician at the Guinness brewery at Dublin
- t-statistic is a signal-to-noise measure: $t = \text{effect} / \text{standard deviation}$
- t-distribution is an approximation to the normal distribution for small samples
- t-contrasts are simply linear combinations of the betas
 \Rightarrow the t-statistic does not depend on the scaling of the regressors or on the scaling of the contrast
- Unilateral test in SPM:

$$H_0 : c^T \beta = 0 \text{ vs. } H_1 : c^T \beta > 0$$



t-contrasts – SPM{t}



Question:

box-car amplitude > 0 ?

$$= H_1 = c^T \beta > 0 ?$$

Null hypothesis:

$$H_0: c^T \beta = 0$$

Test statistic:

*contrast of
parameter
estimates*

$$t = \frac{\text{contrast of parameter estimates}}{\sqrt{\text{variance estimate}}}$$

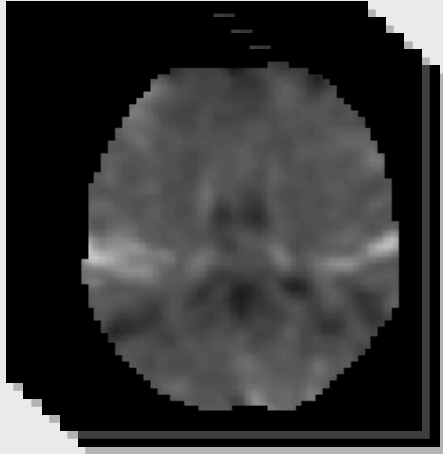


$$p(y | c^T \hat{\beta} = 0)$$

$$t = \frac{c^T \hat{\beta}}{\hat{s} \sqrt{d(c^T \hat{\beta})}} = \frac{c^T \hat{\beta}}{\sqrt{\hat{\sigma}^2 c^T (X^T X)^{-1} c}} \sim t_{N-p}$$

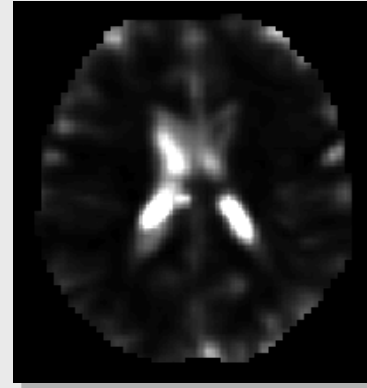
t-contrasts in SPM

For a given contrast c :



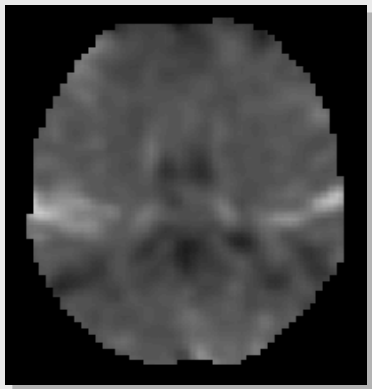
beta_???? images

$$\hat{\beta} = (X^T X)^{-1} X^T y$$



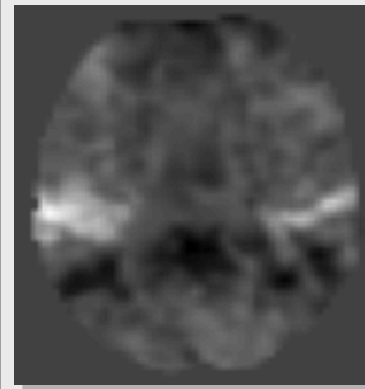
ResMS image

$$\hat{\sigma}^2 = \frac{\hat{e}^T \hat{e}}{N - p}$$



con_???? image

$$c^T \hat{\beta}$$



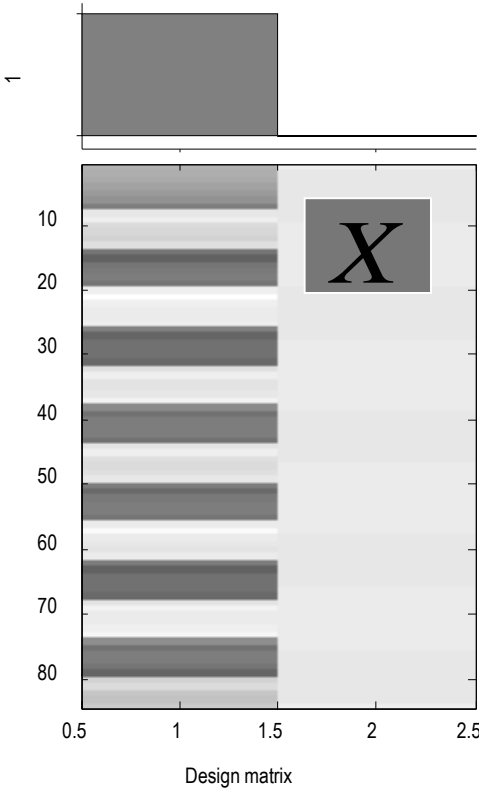
spmT_???? image

SPM{ t }

t-contrast: a simple example

Passive word listening versus rest

$c^T = [1 \quad 0]$

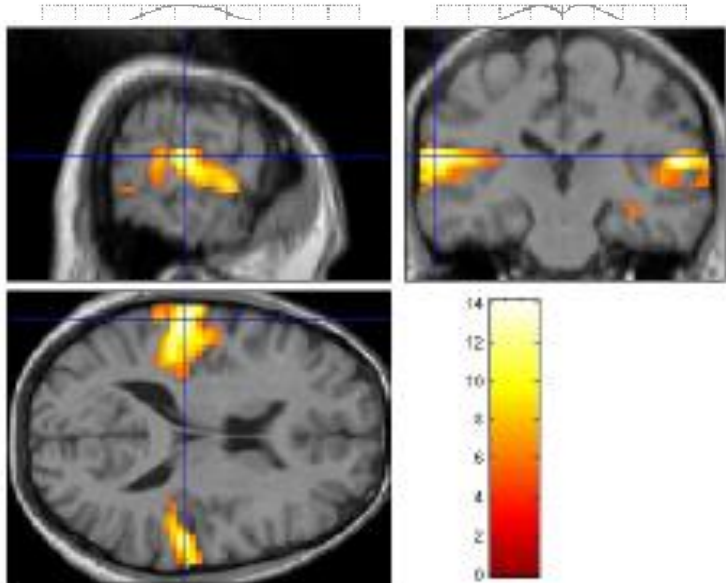


Q: activation during listening ?

Null hypothesis: $\beta_1 = 0$

$$t = \frac{c^T \hat{\beta}}{Std(c^T \hat{\beta})}$$

$$p(y \mid c^T \hat{\beta} = 0)$$



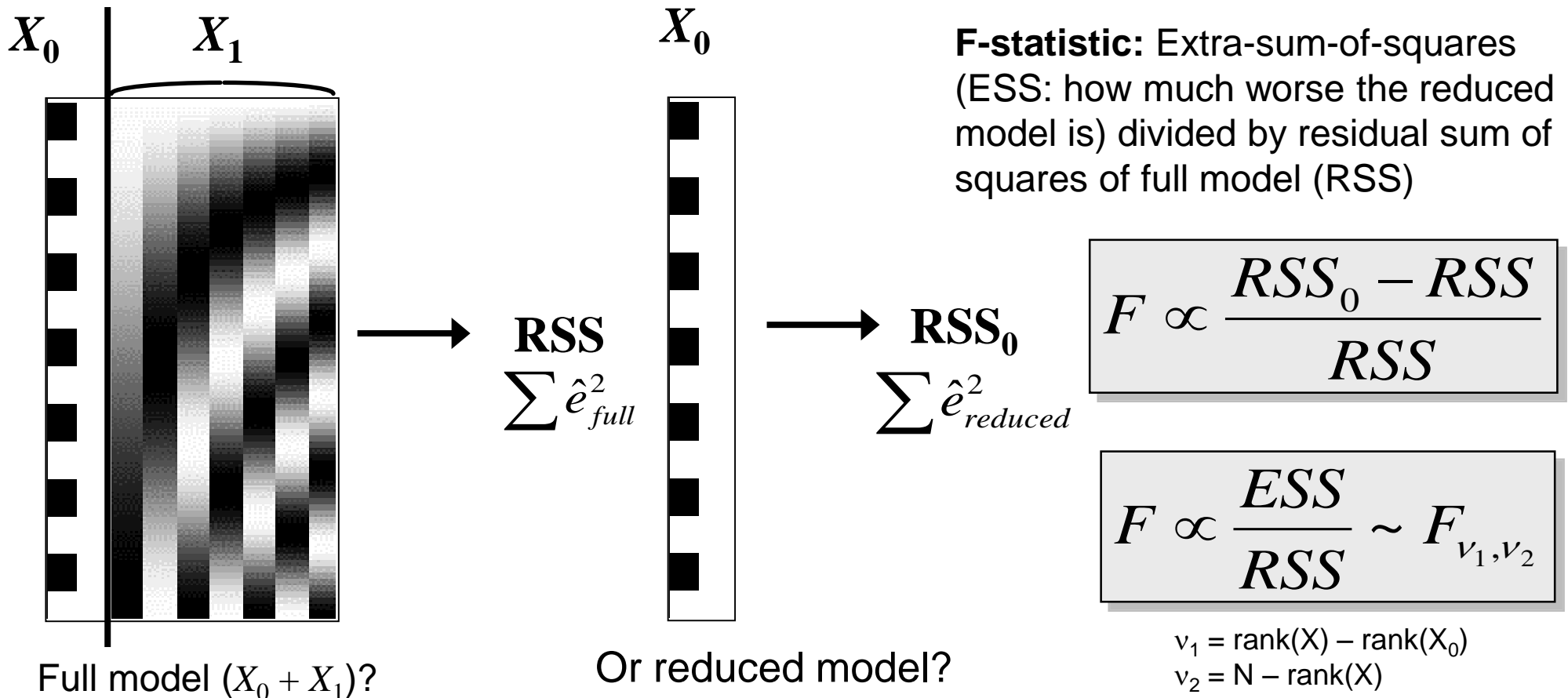
SPMresults:
Height threshold T = 3.2057 {p<0.001}

Statistics: <i>p-values adjusted for search volume</i>										
set-level		cluster-level			voxel-level					mm mm mm
<i>p</i>	<i>c</i>	<i>p</i> corrected	<i>k</i> _E	<i>p</i> uncorrected	<i>p</i> FWE-corr	<i>p</i> FDR-corr	<i>T</i>	(<i>Z</i> _≡)	<i>p</i> uncorrected	
0.000	10	0.000	520	0.000	0.000	0.000	13.94	Inf	0.000	-63 -27 15
					0.000	0.000	12.04	Inf	0.000	-48 -33 12
					0.000	0.000	11.82	Inf	0.000	-66 -21 6
		0.000	426	0.000	0.000	0.000	13.72	Inf	0.000	57 -21 12
					0.000	0.000	12.29	Inf	0.000	63 -12 -3
					0.000	0.000	9.89	7.83	0.000	57 -39 6
		0.000	35	0.000	0.000	0.000	7.39	6.36	0.000	36 -30 -15
		0.000	9	0.000	0.000	0.000	6.84	5.99	0.000	51 0 48
		0.002	3	0.024	0.001	0.000	6.36	5.65	0.000	-63 -54 -3
		0.000	8	0.001	0.001	0.000	6.19	5.53	0.000	-30 -33 -18
		0.000	9	0.000	0.003	0.000	5.96	5.36	0.000	36 -27 9
		0.005	2	0.058	0.004	0.000	5.84	5.27	0.000	-45 42 9
		0.015	1	0.166	0.022	0.000	5.44	4.97	0.000	48 27 24
		0.015	1	0.166	0.036	0.000	5.32	4.87	0.000	36 -27 42

F-test: the extra-sum-of-squares principle

Model comparison: Full vs. reduced model

Null Hypothesis H_0 : True model is X_0 (reduced model)



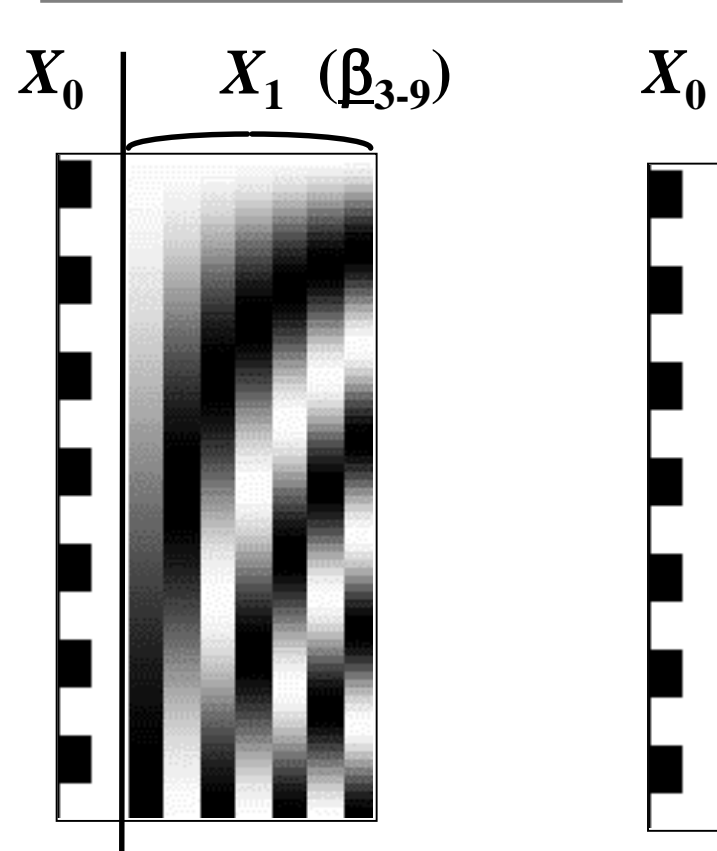
F-test: multidimensional contrasts – SPM{F}

Tests multiple linear hypotheses:

H₀: True model is X_0

$$\mathbf{H}_0: \beta_3 = \beta_4 = \dots = \beta_9 = 0$$

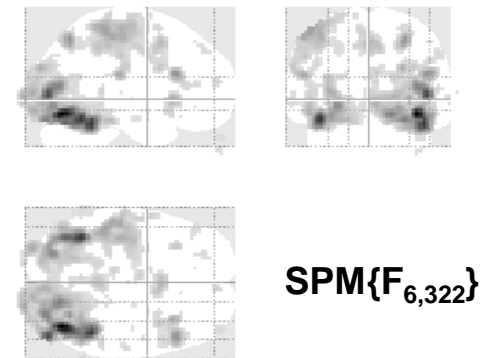
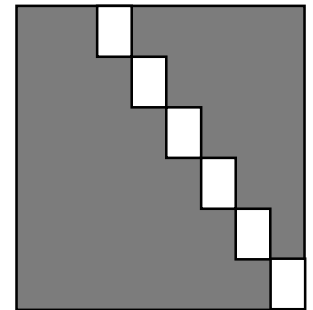
test $H_0: c^T \beta = 0$?



Full model?

Reduced model?

$$c^T = \begin{array}{cccccccc} 0 & 0 & \color{red}{1} & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \color{red}{1} & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & \color{red}{1} & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & \color{red}{1} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & \color{red}{1} & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & \color{red}{1} \end{array}$$



SPM{F_{6,322}}

F-test: a few remarks

- F-tests can be viewed as testing for the additional variance explained by a larger model wrt. a simpler (nested) model \Rightarrow model comparison

- Hypotheses:

$$\begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

Null hypothesis H_0 :

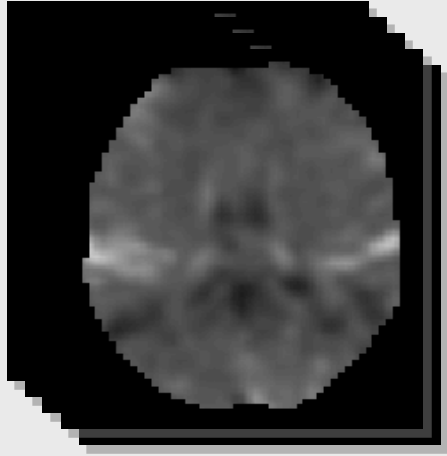
$$\beta_1 = \beta_2 = \dots = \beta_p = 0$$

Alternative hypothesis H_1 :

At least one $\beta_k \neq 0$

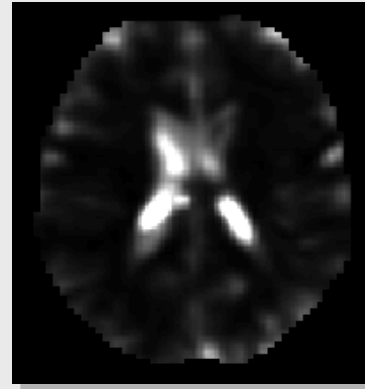
- F-tests are not directional:
When testing a uni-dimensional contrast with an F -test, for example $\beta_1 - \beta_2$, the result will be the same as testing $\beta_2 - \beta_1$.

F-contrast in SPM



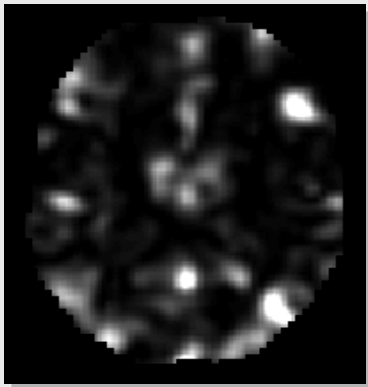
beta_???? images

$$\hat{\beta} = (X^T X)^{-1} X^T y$$



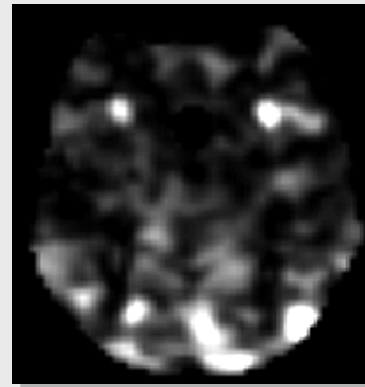
ResMS image

$$\hat{\sigma}^2 = \frac{\hat{e}^T \hat{e}}{N - p}$$



ess_???? images

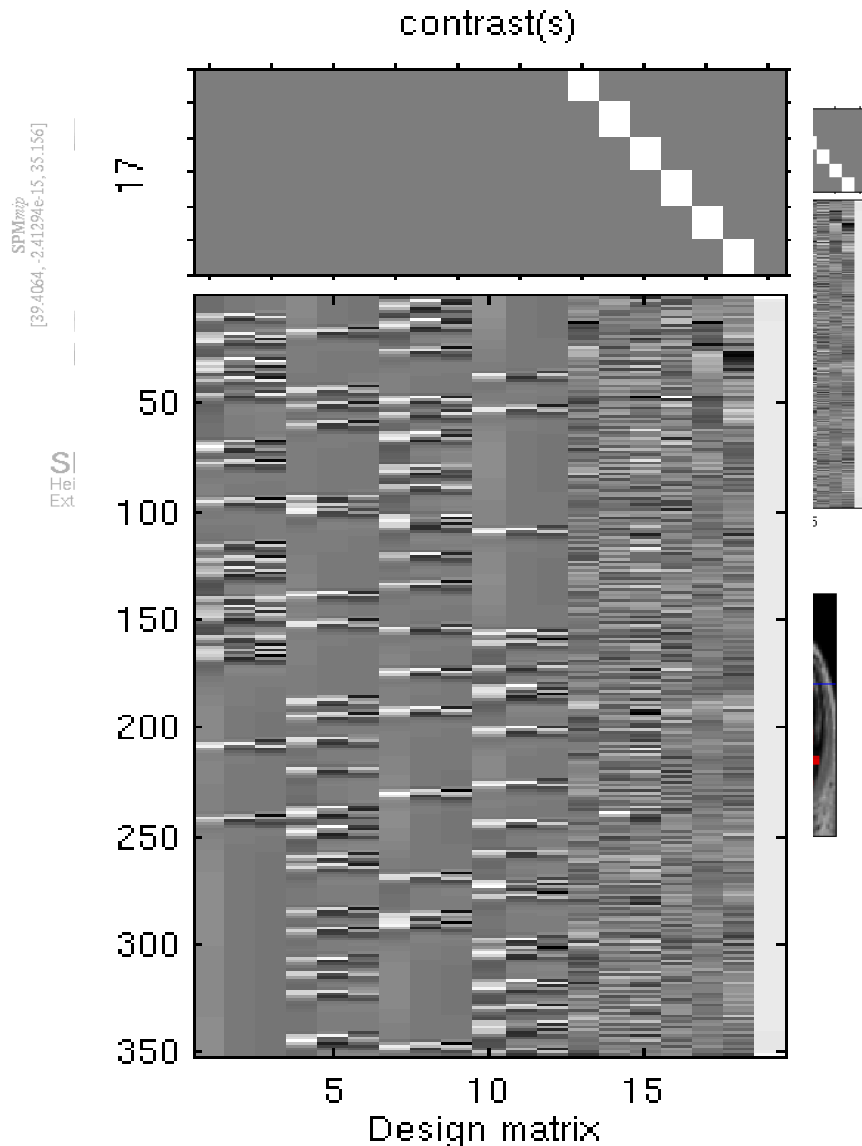
$$(RSS_0 - RSS)$$



spmF_???? images

SPM{F}

F-test example: movement related effects

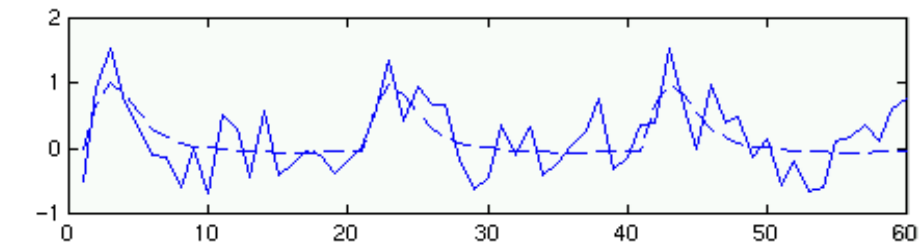


To assess movement-related activation:

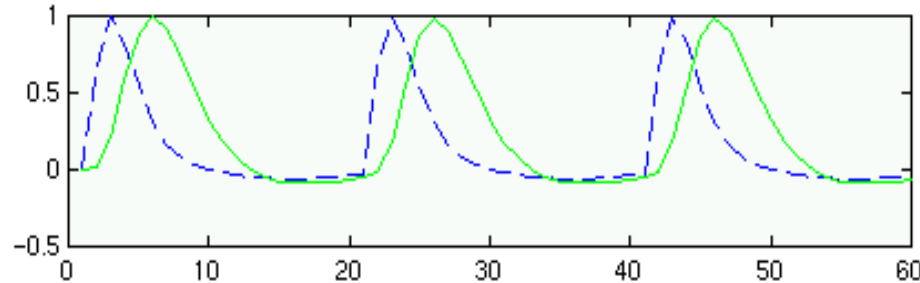
There is a lot of residual movement-related artifact in the data (despite spatial realignment), which tends to be concentrated near the boundaries of tissue types.

By including the realignment parameters in our design matrix, we can “regress out” linear components of subject movement, reducing the residual error, and hence improve our statistics for the effects of interest.

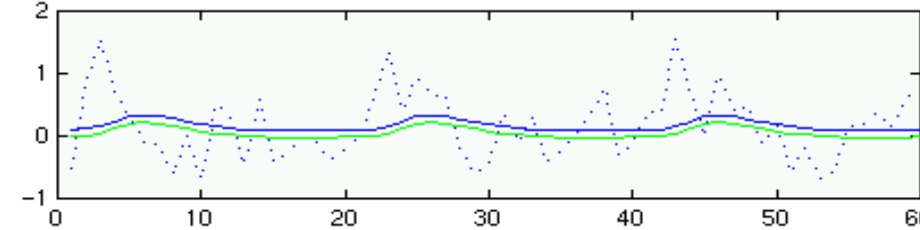
Example: a suboptimal model



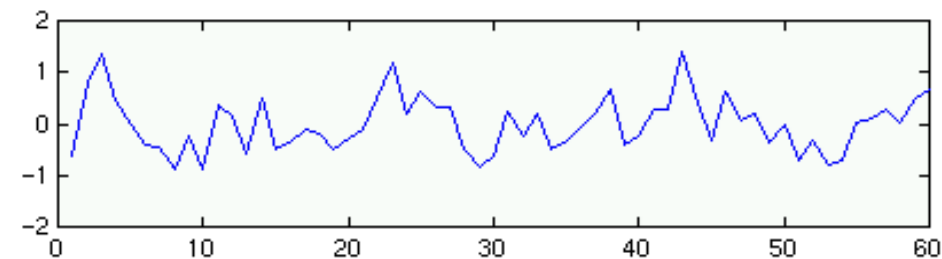
True signal (--) and observed signal



Green regressor is temporally misaligned



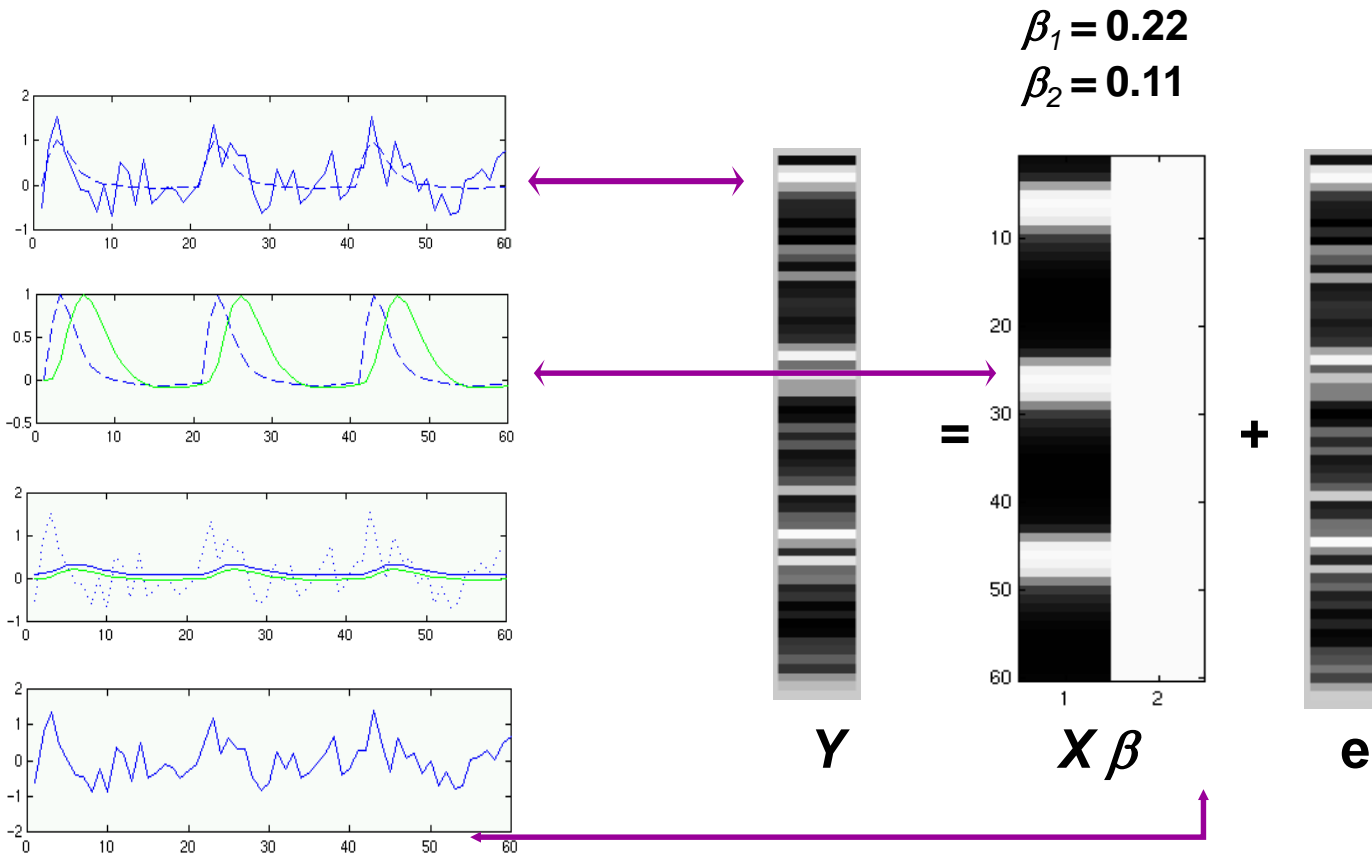
Fitting ($\beta_1 = 0.2$, β_2 (const.) = 0.11);
(here: blue solid line = total fit)



Residuals (still contain some signal)

⇒ Test for the green regressor not significant

Example: a suboptimal model

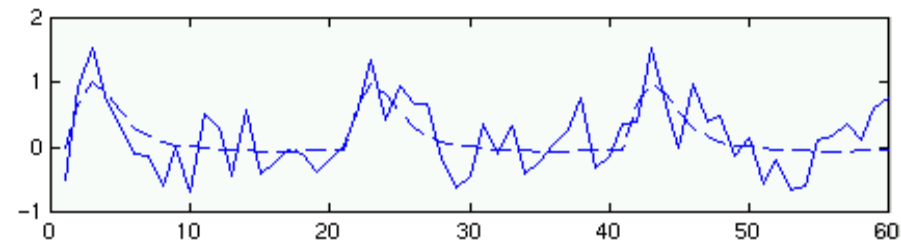


Residual Var. = 0.3

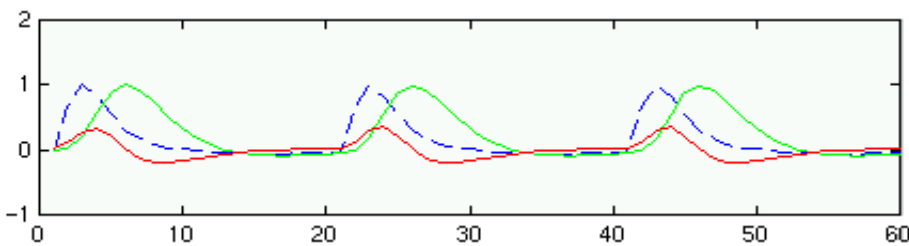
$p(Y/ b_1 = 0) \Rightarrow$
 $p\text{-value} = 0.1$
(t -test)

$p(Y/ b_1 = 0) \Rightarrow$
 $p\text{-value} = 0.2$
(F -test)

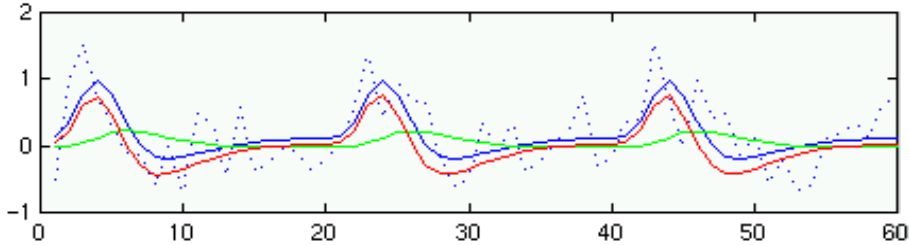
A better model



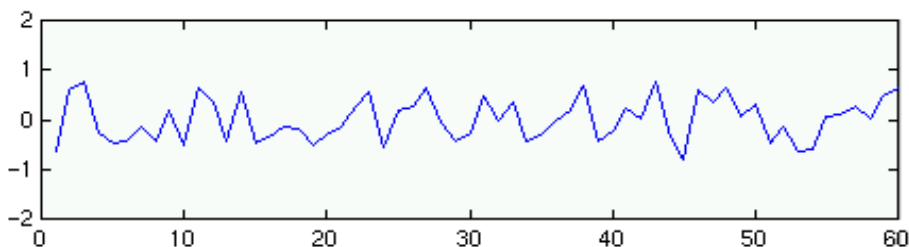
True signal + observed signal



Model (green and red)
and true signal (blue ---)
Red regressor: temporal derivative of
the green regressor



Total fit (blue)
and partial fit (green & red)
Adjusted and fitted signal

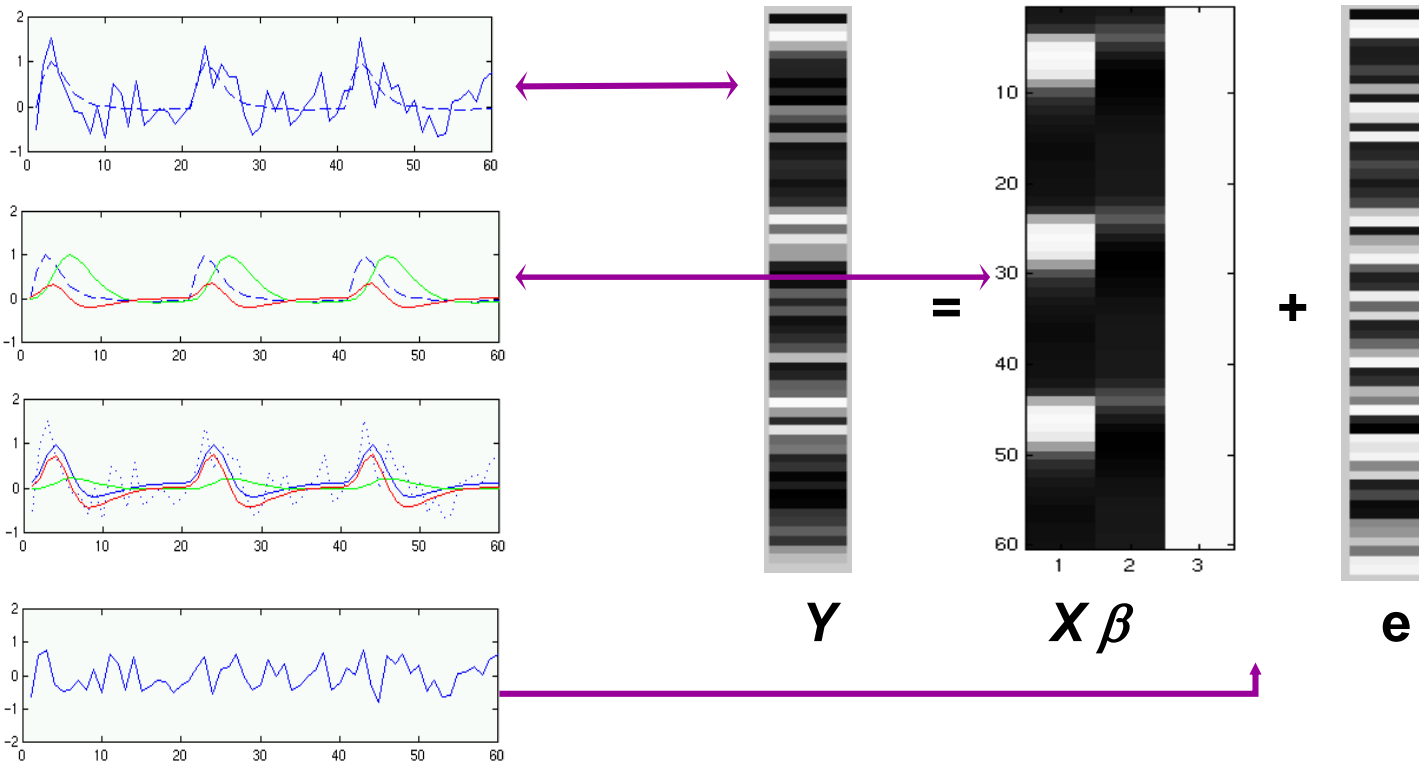


Residuals (less variance & structure)

- ⇒ t -test of the green regressor almost significant
- ⇒ F -test very significant
- ⇒ t -test of the red regressor very significant

A better model

$$\begin{aligned}\beta_1 &= 0.22 \\ \beta_2 &= 2.15 \\ \beta_3 &= 0.11\end{aligned}$$

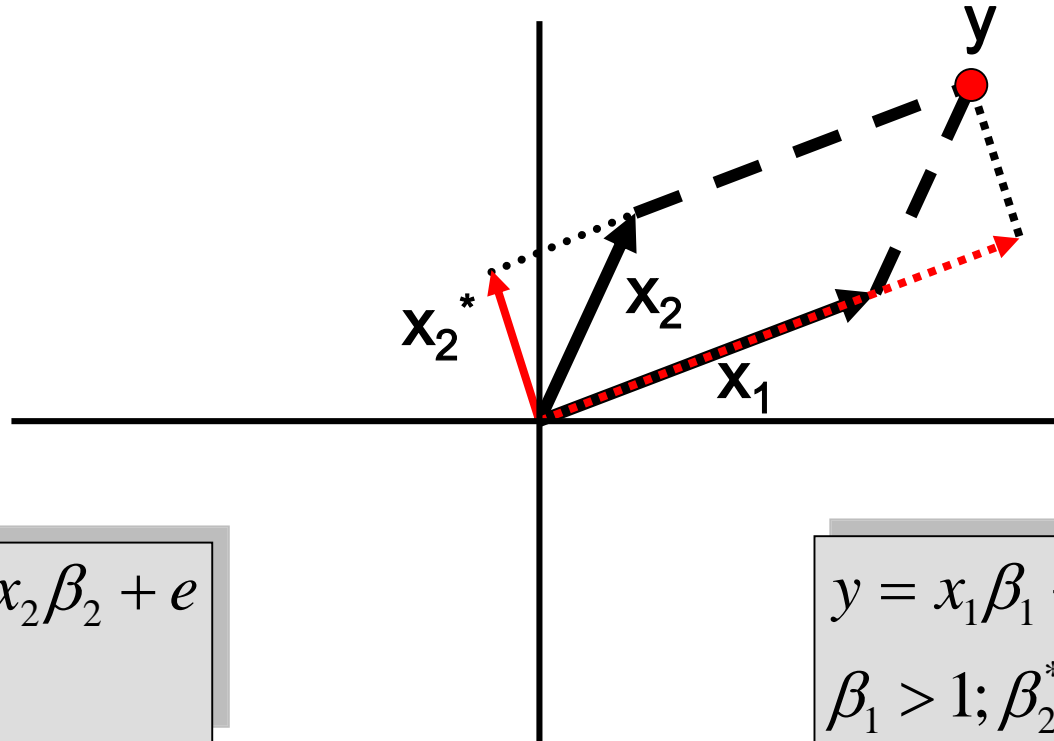


Residual Var. = 0.2

$$\begin{aligned}p(Y/ b_1 = 0) &\Rightarrow \\ p\text{-value} &= 0.07 \\ &\text{(t-test)}\end{aligned}$$

$$\begin{aligned}p(Y/ b_1 = 0, b_2 = 0) &\Rightarrow \\ p\text{-value} &= 0.000001 \\ &\text{(F-test)}\end{aligned}$$

Recap from previous lecture: Correlation among regressors



$$y = x_1\beta_1 + x_2\beta_2 + e$$
$$\beta_1 = \beta_2 = 1$$

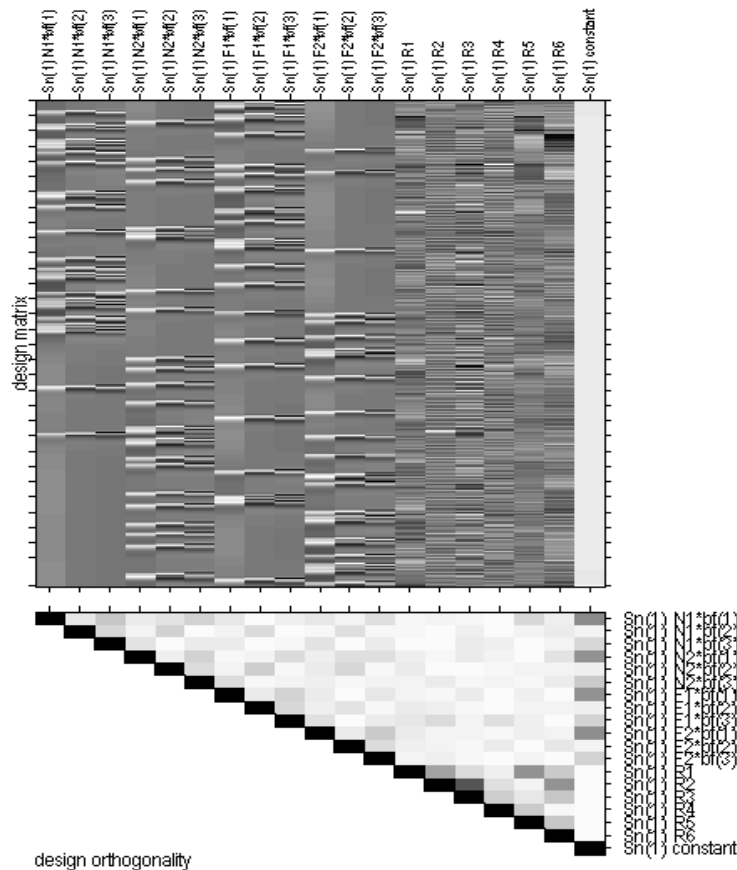
Correlated regressors =
explained variance is shared
between regressors

$$y = x_1\beta_1 + x_2^*\beta_2^* + e$$
$$\beta_1 > 1; \beta_2^* = 1$$

When x_2 is orthogonalized with
regard to x_1 , only the parameter
estimate for x_1 changes, not that
for x_2 !

Design orthogonality

Statistical analysis: Design orthogonality



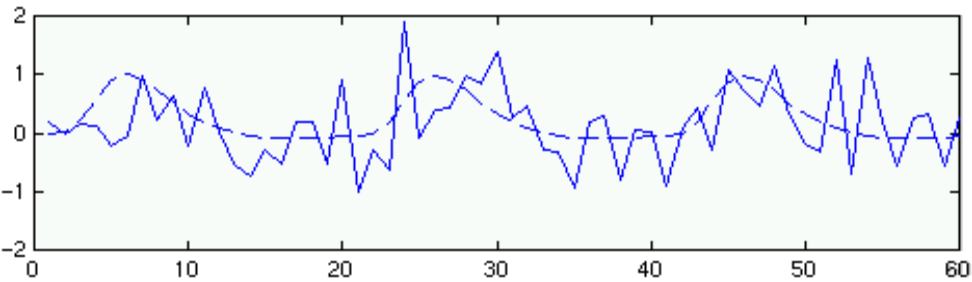
- For each pair of columns of the design matrix, the orthogonality matrix depicts the magnitude of the **cosine of the angle** between them, with the range 0 to 1 mapped from white to black.
- The cosine of the angle between two vectors a and b is obtained by:

$$\cos \alpha = \frac{ab}{|a||b|}$$

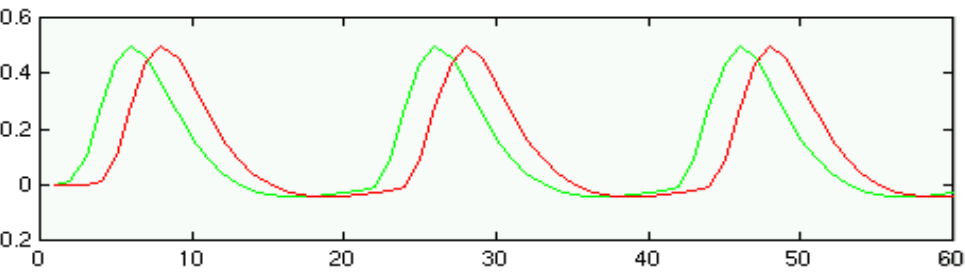
- For **zero-mean vectors**, the cosine of the angle between the vectors is the same as the **correlation** between the two variates:

$$\cos \alpha = \text{corr}_{a,b}$$

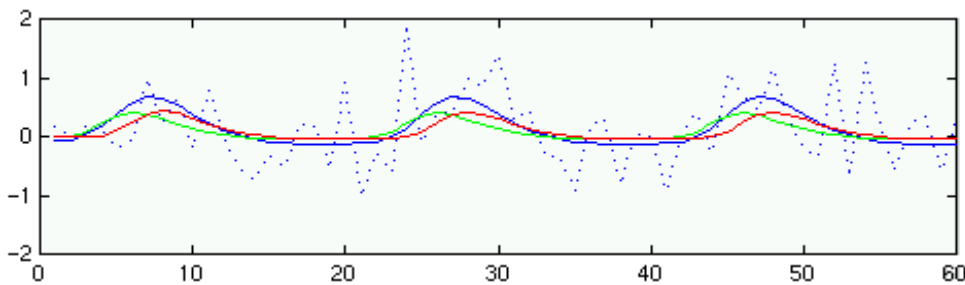
Correlated regressors



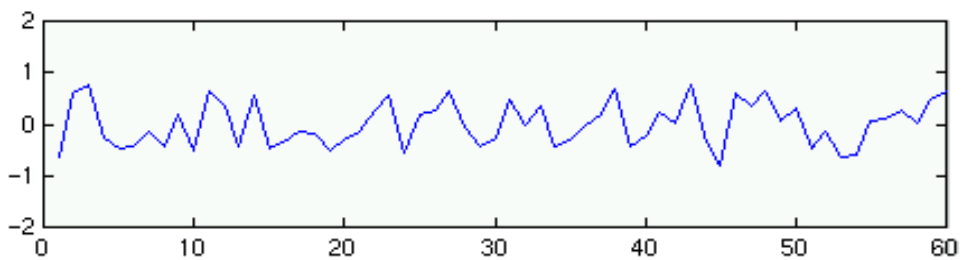
True signal



Model (green and red)

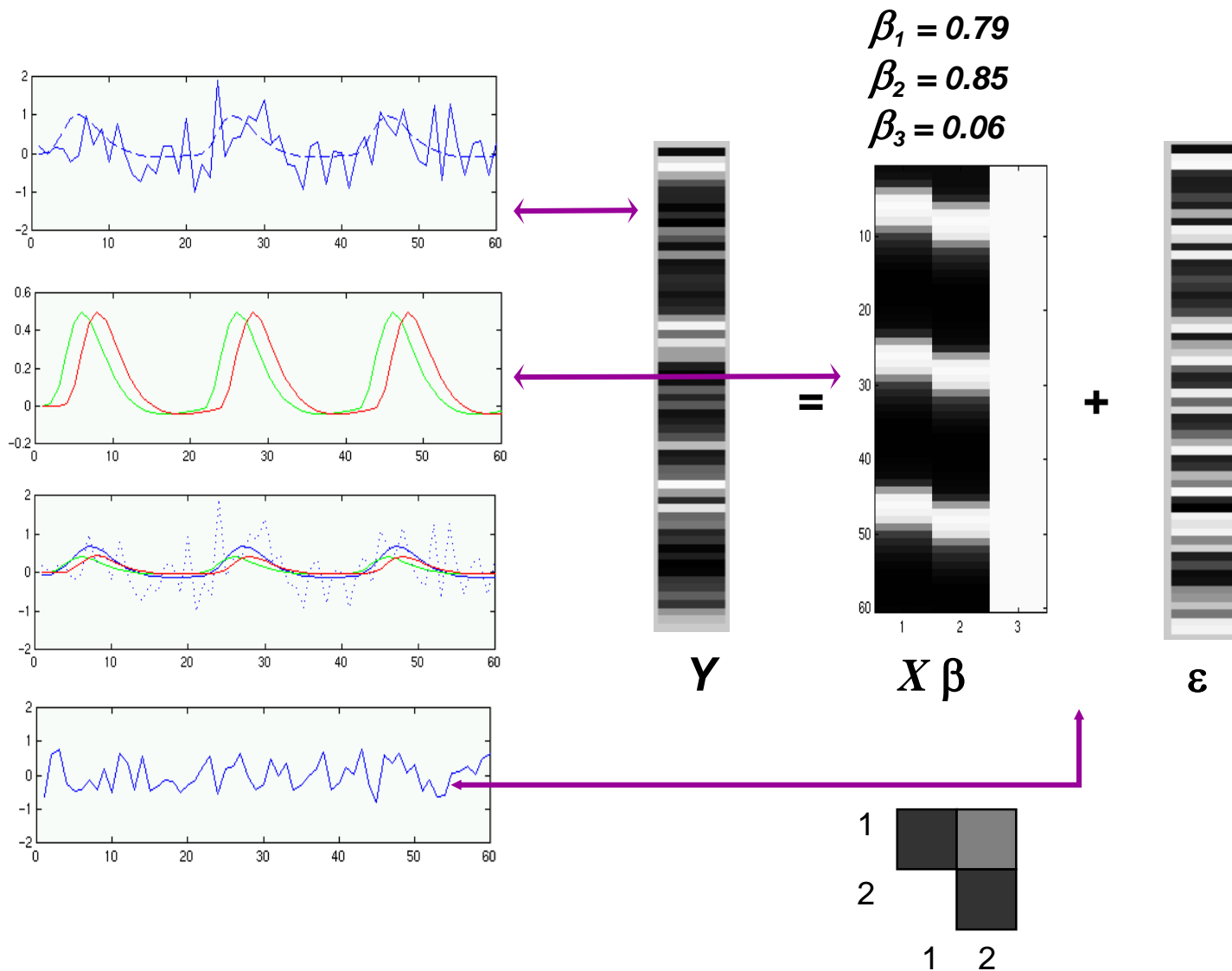


Fit (blue: total fit)



Residual

Correlated regressors



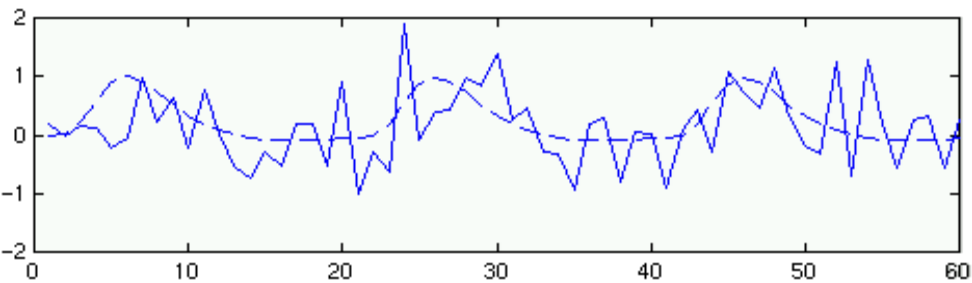
Residual var. = 0.3

$p(Y/\mathbf{b}_1 = 0) \Rightarrow$
 $p\text{-value} = 0.08$
 (t -test)

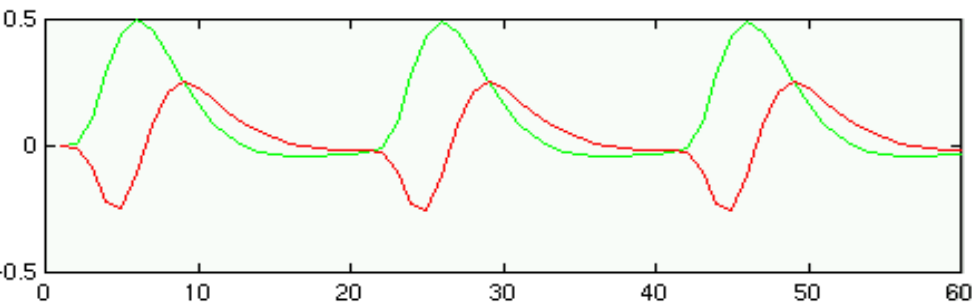
$P(Y/\mathbf{b}_2 = 0) \Rightarrow$
 $p\text{-value} = 0.07$
 (t -test)

$p(Y/\mathbf{b}_1 = 0, \mathbf{b}_2 = 0) \Rightarrow$
 $p\text{-value} = 0.002$
 (F -test)

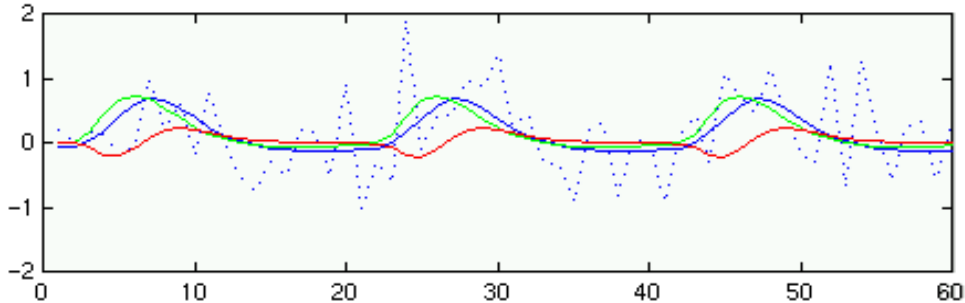
After orthogonalisation



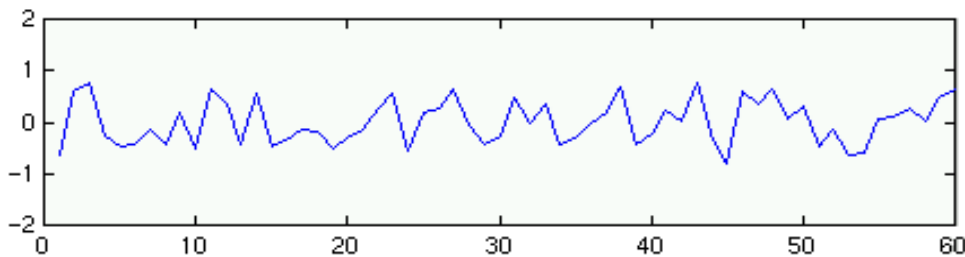
True signal



Model (green and red)
red regressor has been
orthogonalised with respect to the green
one
 \Leftrightarrow remove everything that correlates with
the green regressor

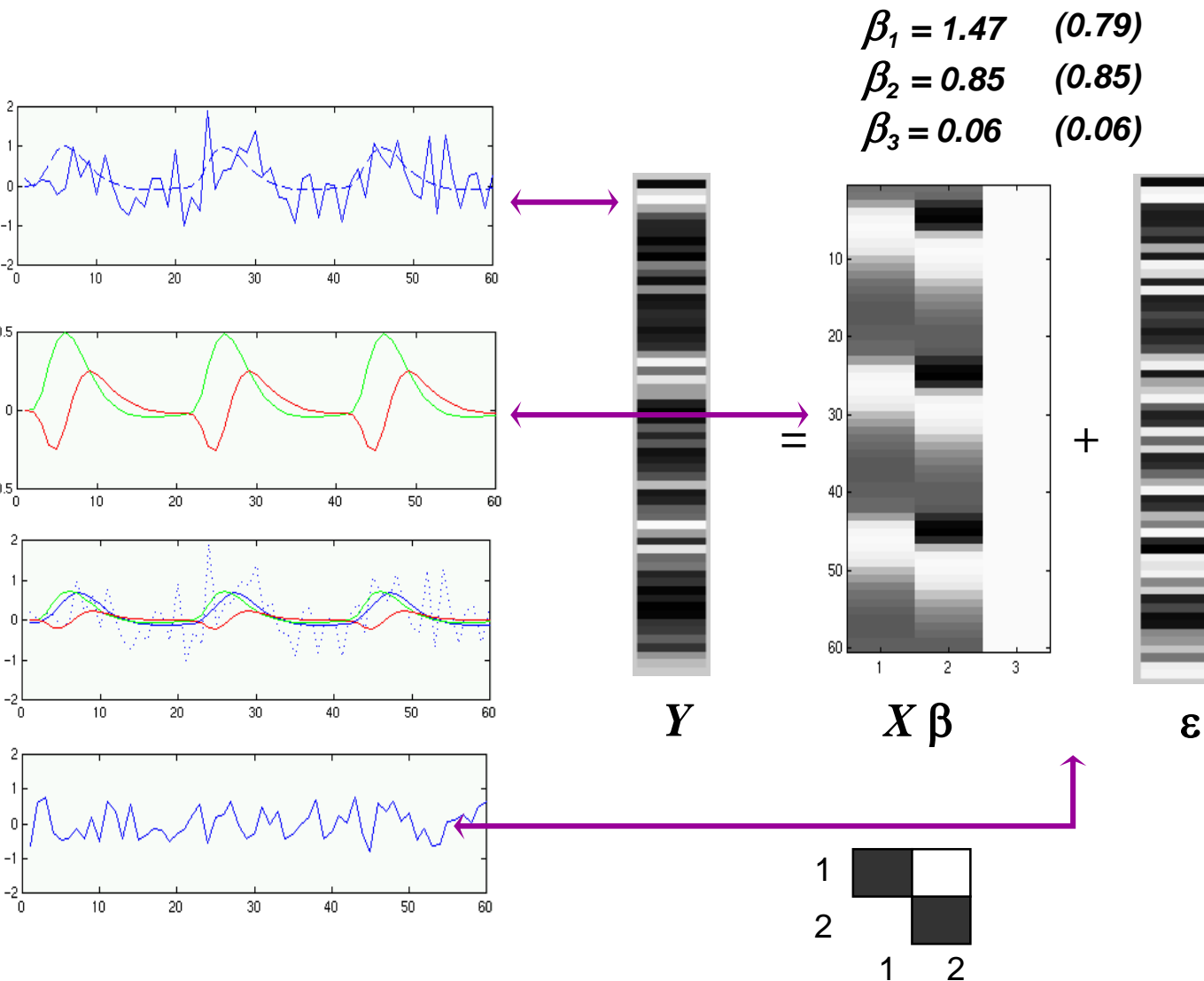


Fit (does not change)



Residuals (do not change)

After orthogonalisation



Residual var. = 0.3

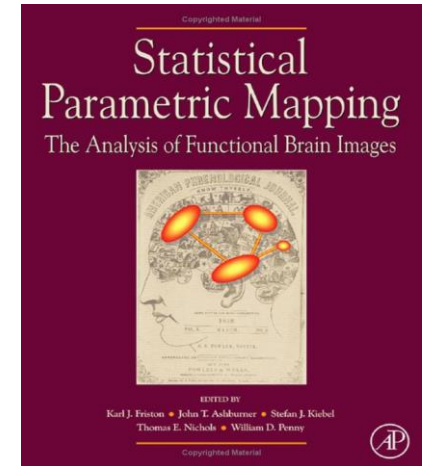
$p(Y | \mathbf{b}_1 = 0)$
 $p\text{-value} = 0.0003$ does change
 (t -test)

$p(Y | \mathbf{b}_2 = 0)$
 $p\text{-value} = 0.07$ does not change
 (t -test)

$p(Y | \mathbf{b}_1 = 0, \mathbf{b}_2 = 0)$
 $p\text{-value} = 0.002$ does not change
 (F -test)

Bibliography

- Friston KJ et al. (2007) *Statistical Parametric Mapping: The Analysis of Functional Brain Images*. Elsevier.



- Christensen R (1996) *Plane Answers to Complex Questions: The Theory of Linear Models*. Springer.
- Friston KJ et al. (1995) Statistical parametric maps in functional imaging: a general linear approach. *Human Brain Mapping* 2: 189-210.

Thank you